Chapter Twentyone.

## Monte Carlo part II: variance reduction by antithetic variates

Outline Solutions to odd-numbered exercises from the book: An Introduction to Financial Option Valuation: Mathematics, Stochastics and Computation, by Desmond J. Higham, Cambridge University Press, 2004 ISBN 0521 83884 3 (hardback) ISBN 0521 54757 1 (paperback)

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**21.1** Starting from (21.1) and using (3.6) and (3.7), we have

$$\begin{aligned} \mathsf{cov}(X,Y) &= & \mathbb{E}\left[(X - \mathbb{E}(X))(Y - \mathbb{E}(Y))\right] \\ &= & \mathbb{E}\left[XY - X\mathbb{E}(Y) - \mathbb{E}(X)Y + \mathbb{E}(X)\mathbb{E}(Y))\right] \\ &= & \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) - \mathbb{E}(X)\mathbb{E}(Y) + \mathbb{E}(X)\mathbb{E}(Y) \\ &= & \mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y). \end{aligned}$$

If X ad Y are independent then  $\mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y) = 0$  (see (3.9)), so  $\operatorname{cov}(X, Y) = 0$ .

**21.3** Using (3.6) and (3.10), we have

$$\begin{split} \mathbb{V}ar(X+Y) &= \mathbb{E}((X+Y)^2) - (\mathbb{E}(X+Y))^2 \\ &= \mathbb{E}((X+Y)^2) - (\mathbb{E}(X) + \mathbb{E}(Y))^2 \\ &= \mathbb{E}(X^2 + 2XY + Y^2) - \mathbb{E}(X)^2 - 2\mathbb{E}(X)\mathbb{E}(Y) - \mathbb{E}(Y)^2 \\ &= \mathbb{E}(X^2) + 2\mathbb{E}(XY) + \mathbb{E}(Y^2) - \mathbb{E}(X)^2 - 2\mathbb{E}(X)\mathbb{E}(Y) - \mathbb{E}(Y)^2 \\ &= \mathbb{E}(X^2) - \mathbb{E}(X)^2 + \mathbb{E}(Y^2) - \mathbb{E}(Y)^2 + 2(\mathbb{E}(XY) - \mathbb{E}(X)\mathbb{E}(Y)) \\ &= \mathbb{V}ar(X) + \mathbb{V}ar(Y) + 2\mathsf{cov}(X,Y). \end{split}$$

- **21.5** Consider the circumstance where f and g are both monotonic increasing. We look at two cases.
  - **Case 1:** If  $x \ge y$  then  $f(x) f(y) \ge 0$  and  $g(x) g(y) \ge 0$ , so  $(f(x) f(y))(g(x) g(y)) \ge 0$ .
  - **Case 2:** If  $y \ge x$  then  $f(x) f(y) \le 0$  and  $g(x) g(y) \le 0$ , so  $(f(x) f(y))(g(x) g(y)) \ge 0$ . (We could also just refer to Case 1 and say "by symmetry".)

A similar argument works for the monotonic decreasing case.

**21.7** For this f, when  $U \sim U(0, 1)$  we have  $f(U_i) + f(1-U_i) = \alpha U_1 + \beta + \alpha(1-U_i) + \beta = \alpha + 2\beta$ , so the sum in (21.13) reduces to  $\frac{1}{2}\alpha + \beta$  (which is non-random). The true expected value  $\mathbb{E}(f(U))$  is  $\mathbb{E}(\alpha U + \beta) = \alpha \mathbb{E}(U) + \beta = \frac{1}{2}\alpha + \beta$ . Hence the estimate (21.13) is exact for this f. Each computed sample of  $\frac{1}{2}(f(U_i) + f(1 - U_i))$  has the same, non-random, value, so the sample variance is zero. The confidence interval then collapses to the point  $\frac{1}{2}\alpha + \beta$ ; this correctly reflects the fact that the sample mean is exact.

Similarly, when  $U \sim \mathsf{N}(0, 1)$  we have  $f(U_i) + f(-U_i) = \alpha U_1 + \beta + \alpha(-U_i) + \beta = 2\beta$ , so the sum in (21.19) reduces to  $\beta$  (which is non-random). The true expected value  $\mathbb{E}(f(U))$  is  $\mathbb{E}(\alpha U + \beta) = \alpha \mathbb{E}(U) + \beta = \beta$ . Hence the estimate (21.19) is exact for this f. Each computed sample of  $\frac{1}{2}(f(U_i) + f(-U_i))$  has the same, non-random, value, so the sample variance is zero. The confidence interval then collapses to the point  $\beta$ ; this correctly reflects the fact that the sample mean is exact.