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Royal Netherlands
Meteorological Institute
Ministry of Infrastructure and the
Environment

Computation for very large multiscale spatio-temporal conditional distributions

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THE UNIVERSITY of EDINBURGH

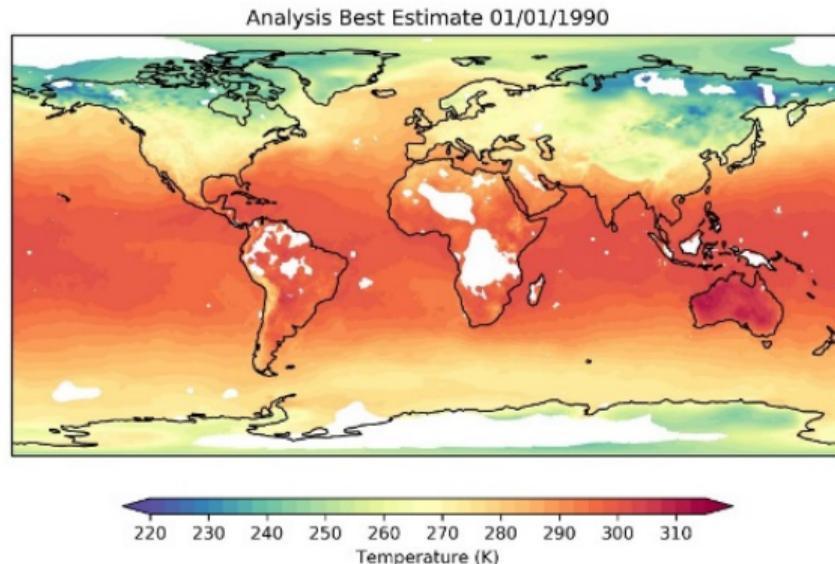
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EUSTACE ANALYSIS

Combines in-situ and satellite data sources to derive daily air temperatures across the globe with quantified uncertainties.

- Daily mean air temperature (2 m) estimates from the mid-late 19th century at $\frac{1}{4}$ degree resolution.
- Observational dataset for use in climate monitoring, services and research.
 - Quantify bias and uncertainty arising from observational sampling (in space and time);
 - Quantify uncertainty from instrumental effects/network changes.
- Higher resolution daily gridded analyses for regional climate
 - Combine in situ and remote sensing data to support high resolution analysis.
 - Absolute temperature rather than anomaly product.



Standardised observation uncertainty models

- ▶ Each data source may have complicated dependence structure
- ▶ To facilitate information blending, use a common error term structure

Common satellite derived data error model framework

The observational&calibration errors are modelled as three error components:

- ▶ independent (ϵ_0),
- ▶ spatially and/or temporally correlated (ϵ_1), and
- ▶ systematic (ϵ_2),

with distributions determined by the uncertainty information from satellite calibration models.

$$\text{E.g., } y_i = T_m(\mathbf{s}_i, t_i) + \epsilon_0(\mathbf{s}_i, t_i) + \epsilon_1(\mathbf{s}_i, t_i) + \epsilon_2(\mathbf{s}_i, t_i)$$

In practice, each data source might have several different components of each type; independent components can be merged, but not necessarily correlated or systematic components.

Station observation&homogenisation model

Daily means

For station k at day t_i ,

$$y_m^{k,i} = T_m(\mathbf{s}_k, t_i) + \sum_{j=1}^{J_k} H_j^k(t_i) e_m^{k,j} + \epsilon_m^{k,i},$$

where $H_j^k(t)$ are temporal step functions, $e_m^{k,j}$ are latent bias variables, and $\epsilon_m^{k,i}$ are independent measurement and discretisation errors.

Daily mean/max/min

For station k at day t_i ,

$$y_m^{k,i} = T_m(\mathbf{s}_k, t_i) + \tilde{H}_m^k(t_i) + \epsilon_m^{k,i},$$

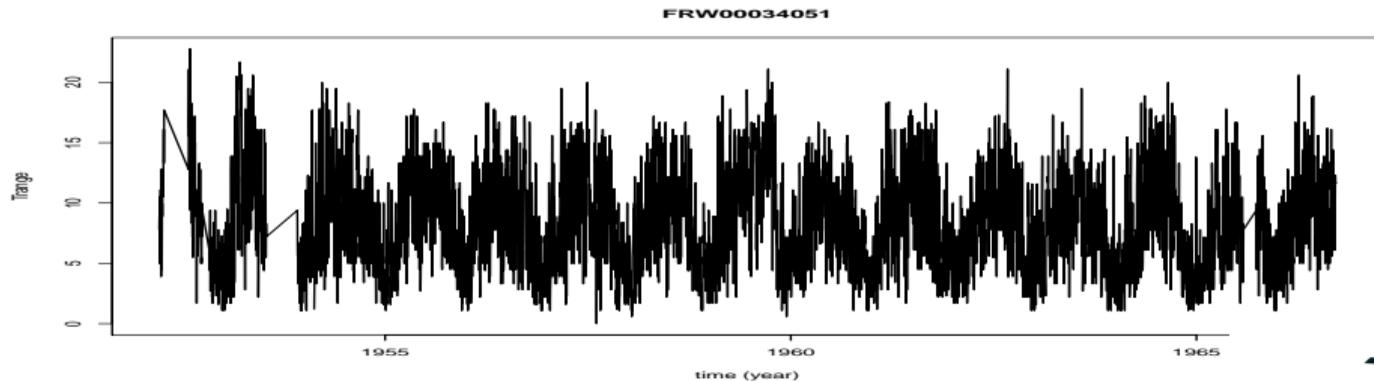
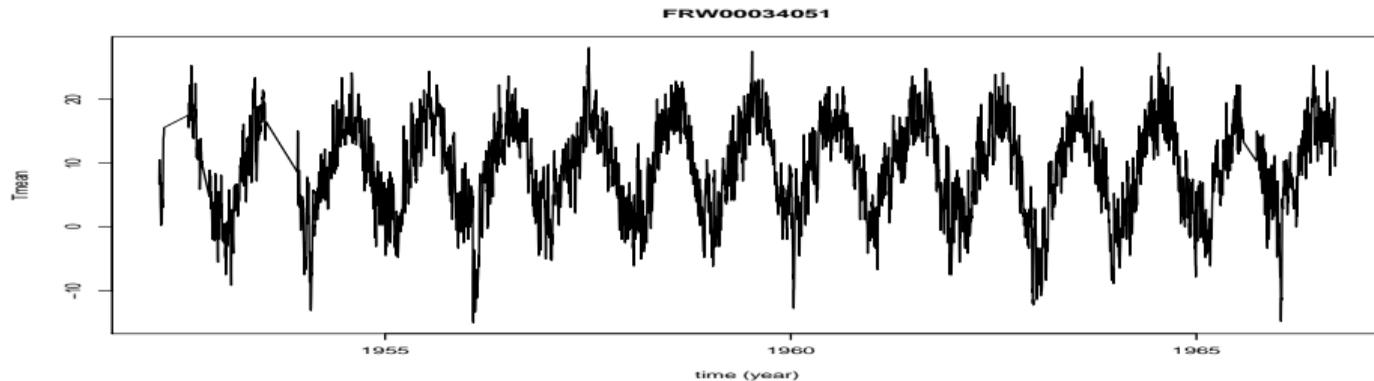
$$y_x^{k,i} = T_m(\mathbf{s}_k, t_i) + \frac{\exp[\tilde{H}_r^k(t_i)]}{2} T_r(\mathbf{s}_k, t_i) + \epsilon_x^{k,i},$$

$$y_n^{k,i} = T_m(\mathbf{s}_k, t_i) - \frac{\exp[\tilde{H}_r^k(t_i)]}{2} T_r(\mathbf{s}_k, t_i) + \epsilon_n^{k,i},$$

where \tilde{H}_\cdot are the total bias correction variables for each observation.

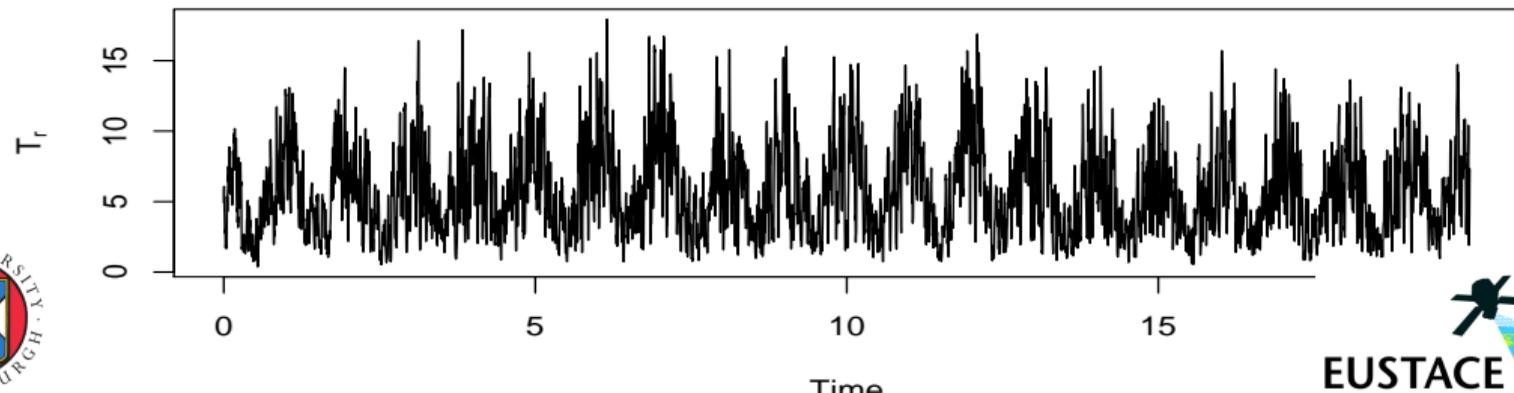
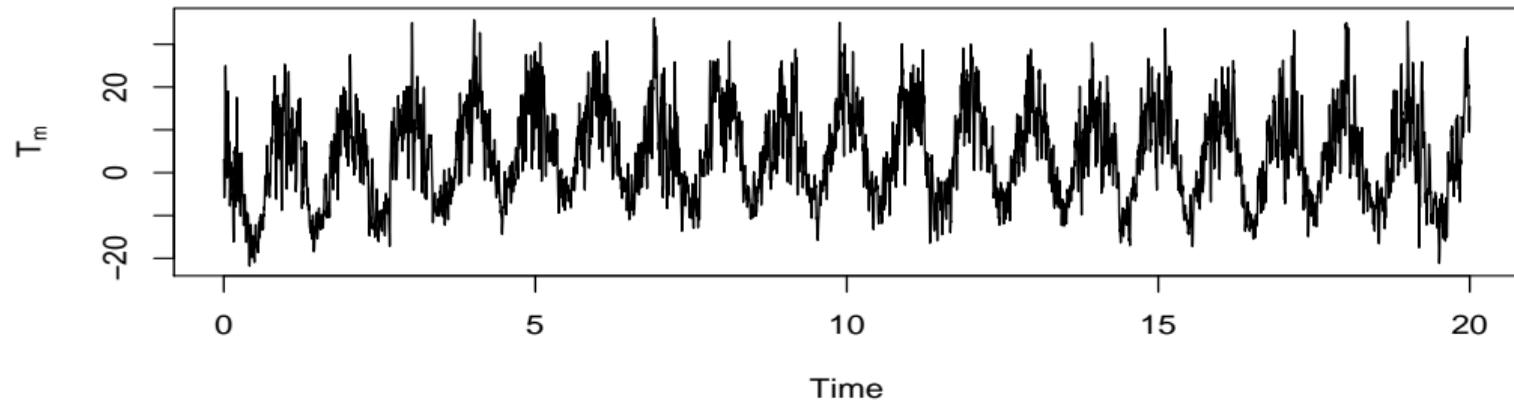
Observed data

Observed daily T_{mean} and T_{range} for station FRW00034051

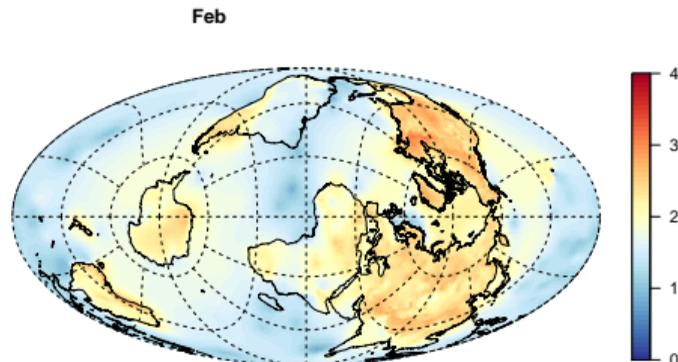
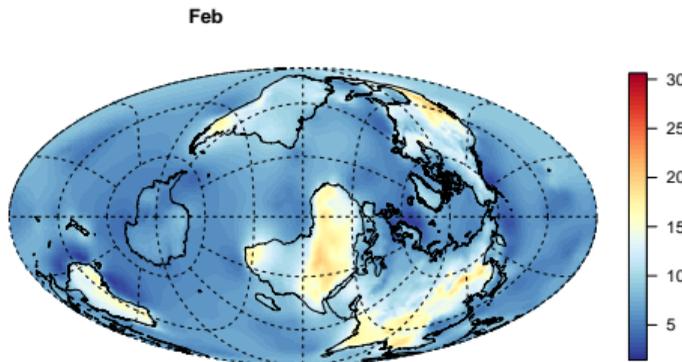
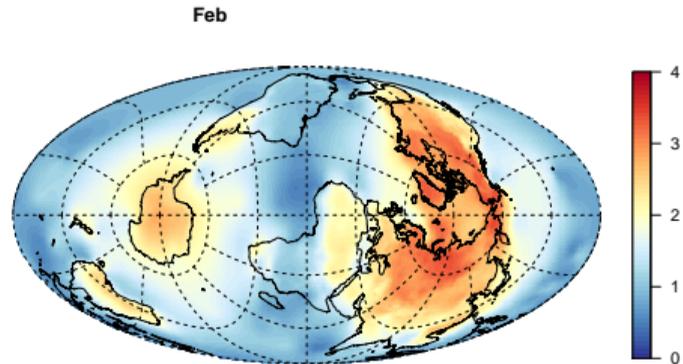
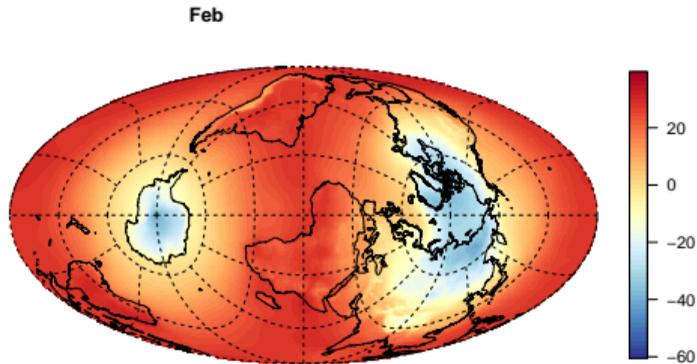


Combined model samples for T_m and T_r

(Proof of concept; no actual data was involved in this figure)



Estimates of median & scale for T_m and T_r



February climatology

(Preliminary estimates, using only in-situ land station data)

Matérn driven heat equation on the sphere

The iterated heat equation is a simple non-separable space-time SPDE family:

$$(\kappa^2 - \Delta)^{\gamma/2} \left[\phi \frac{\partial}{\partial t} + (\kappa^2 - \Delta)^{\alpha/2} \right]^\beta x(\mathbf{s}, t) = \dot{W}(\mathbf{s}, t) / \tau$$

For constant parameters, $x(\mathbf{s}, t)$ has spatial Matérn covariance (for each t).

Discrete domain Gaussian Markov random fields (GMRFs)

$\mathbf{x} = (x_1, \dots, x_n) \sim \mathcal{N}(\boldsymbol{\mu}, \mathbf{Q}^{-1})$ is Markov with respect to a neighbourhood structure $\{\mathcal{N}_i, i = 1, \dots, n\}$ if $Q_{ij} = 0$ whenever $j \notin \mathcal{N}_i \cup i$.

- Project the SPDE solution space onto local basis functions:
random Markov dependent basis weights (Lindgren et al, 2011).

A finite element approximation has structure

$$x(\mathbf{s}, t) = \sum_{i,j} \psi_i^{[s]}(\mathbf{s}) \psi_j^{[t]}(t) x_{ij}, \quad \mathbf{x} \sim \mathcal{N}(\mathbf{0}, \mathbf{Q}^{-1}), \quad \mathbf{Q} = \sum_{k=0}^{\alpha+\beta+\gamma} \mathbf{Q}_{t,k} \otimes \mathbf{Q}_{s,k}$$

even, e.g., if the spatial scale parameter κ is spatially varying.

Linearised inference

All Spatio-temporal latent random processes combined into $\mathbf{x} = (\mathbf{u}, \boldsymbol{\beta}, \mathbf{b})$, with joint expectation $\boldsymbol{\mu}_x$ and precision \mathbf{Q}_x :

$$(\mathbf{x} \mid \boldsymbol{\theta}) \sim \mathcal{N}(\boldsymbol{\mu}_x, \mathbf{Q}_x^{-1}) \quad (\text{Prior})$$

$$(\mathbf{y} \mid \mathbf{x}, \boldsymbol{\theta}) \sim \mathcal{N}(h(\mathbf{x}), \mathbf{Q}_{y|x}^{-1}) \quad (\text{Observations})$$

$$p(\mathbf{x} \mid \mathbf{y}, \boldsymbol{\theta}) \propto p(\mathbf{x} \mid \boldsymbol{\theta}) p(\mathbf{y} \mid \mathbf{x}, \boldsymbol{\theta}) \quad (\text{Conditional posterior})$$

Non-linear and/or non-Gaussian observations

For a non-linear $h(\mathbf{x})$ with Jacobian \mathbf{J} at $\mathbf{x} = \tilde{\boldsymbol{\mu}}$, iterate:

$$(\mathbf{x} \mid \mathbf{y}, \boldsymbol{\theta}) \stackrel{\text{approx}}{\sim} \mathcal{N}(\tilde{\boldsymbol{\mu}}, \tilde{\mathbf{Q}}^{-1}) \quad (\text{Approximate conditional posterior})$$

$$\tilde{\mathbf{Q}} = \mathbf{Q}_x + \mathbf{J}^\top \mathbf{Q}_{y|x} \mathbf{J}$$

$$\tilde{\boldsymbol{\mu}}' = \tilde{\boldsymbol{\mu}} + a \tilde{\mathbf{Q}}^{-1} \left\{ \mathbf{J}^\top \mathbf{Q}_{y|x} [\mathbf{y} - h(\tilde{\boldsymbol{\mu}})] - \mathbf{Q}_x (\tilde{\boldsymbol{\mu}} - \boldsymbol{\mu}_x) \right\}$$

for some $a > 0$ chosen by line-search.

Multiscale precision structure

Two-level model with coarse scale x_1 and fine scale $x_0 = z_0 + Bx_1$, observations linked linearly to the fine scale only, $y = Ax_0$.

A priori independent blocks

Blocks (z_0, x_1) , $J = [A \ AB]$:

$$Q_{z_0, x_1 | y} = \begin{bmatrix} Q_0 + A^\top Q_{y|x} A & A^\top Q_{y|x} AB \\ B^\top A^\top Q_{y|x} A & Q_1 + B^\top A^\top Q_{y|x} AB \end{bmatrix}$$

Accumulative blocks

Blocks (x_0, x_1) , $J = [A \ 0]$.

$$Q_{x_0, x_1 | y} = \begin{bmatrix} Q_0 + A^\top Q_{y|x} A & -Q_0 B \\ -B^\top Q_0 & Q_1 + B^\top Q_0 B \end{bmatrix}$$

Multiscale Schur complement approximation

Solving $Q_{x|y}x = b$ can be formulated using two solves with the upper (fine) block $Q_0 + A^\top Q_{y|x}A$, and one solve with the *Schur complement*

$$Q_1 + B^\top Q_0 B - B^\top Q_0 (Q_0 + A^\top Q_{y|x}A)^{-1} Q_0$$

By mapping the fine scale model onto the coarse basis used for the coarse model, we get an *approximate* (and sparse) Schur solve via

$$\begin{bmatrix} \tilde{Q}_B + B^\top A^\top Q_{y|x} A B & -\tilde{Q}_B \\ -\tilde{Q}_B & Q_1 + \tilde{Q}_B \end{bmatrix} \begin{bmatrix} \text{ignored} \\ x_1 \end{bmatrix} = \begin{bmatrix} 0 \\ \tilde{b} \end{bmatrix}$$

where $\tilde{Q}_B = B^\top Q_0 B$.

The block matrix can be interpreted as the precision of a bivariate field on a common, coarse spatio-temporal scale, and the same technique applied to this system, with $x_{1,1} = B_{1|2}x_{1,2} + \text{finer scale variability}$.



Problem: Requires reading the observation metadata multiple times for matrix-free iterations.



Iterative solutions for $\sim 10^{11}$ latent variables

- ▶ Nonlinear Newton iteration with robust line-search
- ▶ Preconditioned conjugate gradient (PCG) iteration for

$$Q(\mu - \hat{\mu}) = r = b - Q\hat{\mu}$$

- ▶ Approximate posterior sampling with PCG: $Q(x - \hat{\mu}) = Lw$
Requires only a rectangular pseudo-Cholesky factorisation $LL^\top = Q$.

Possible due to the kronecker product sum precision structure. Simplified example:

$$\begin{aligned} Q_0 &= Q_{t,1} \otimes Q_{s,1} + Q_{t,2} \otimes Q_{s,2}, \\ Q_{t,k} &= L_{t,k} L_{t,k}^\top, \quad Q_{s,k} = L_{s,k} L_{s,k}^\top, \quad Q_1 = L_1 L_1^\top, \quad Q_{y|x} = L_{y|x} L_{y|x}^\top, \\ L_{x_0, x_1 | y} &= \begin{bmatrix} L_{t,1} \otimes L_{s,1} & L_{t,2} \otimes L_{s,2} & 0 & A^\top L_{y|x} \\ -B^\top (L_{t,1} \otimes L_{s,1}) & -B^\top (L_{t,2} \otimes L_{s,2}) & L_1 & 0 \end{bmatrix} \end{aligned}$$

- ▶ Current implementation uses multiscale-blockwise Gauss-Seidel with temporally independent fine scale
- ▶ Ongoing work on overlapping space-time block preconditioning within each level for temporally dependent fine scale

MULTI-SCALE ANALYSIS MODEL

Statistical model for temperature variations and different scales (space and time):

- **Climatological variation:** local seasonal cycle with effects of latitude, altitude and coastal influence.
- **Large-scale variation:** Slowly varying climatological mean temperature field. Station homogenisation.
- **Daily Local:** daily variability associated with weather. Satellite retrieval biases.

Simultaneously estimates observational biases of known bias structures:

- e.g. satellite biases, station homogenisation.

Processed on STFC's LOTUS cluster www.jasmin.ac.uk:

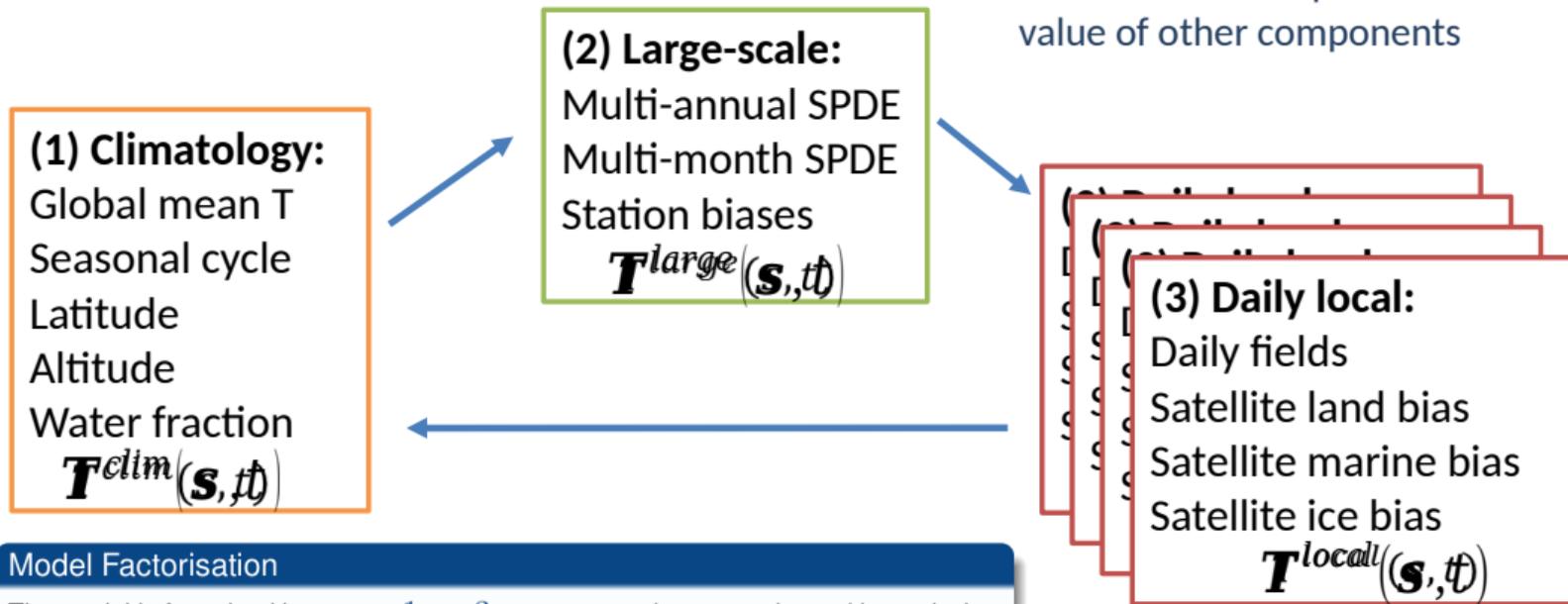
- Largest solves processed on 20 core/256GB RAM node.
- Highly parallel observation pre-processing.

Element	Resolution	N Variables
Seasonal	Bimonthly x 1° SPDE	245,772
Slow-scale*	5 year x 5° SPDE	107,604
Latitude	0.5° latitude SPDE	721
Altitude	(0.25° grid)	1
Coastal	(0.25° grid)	1
Grand mean	Analysis mean	1

Element	Resolution	N Variables
Large-scale	3 monthly x 5° SPDE	1,752,408
Station bias	NA	82,072

Element	Resolution	N Variables per day
Daily local	~0.5 degree SPDE	162,842
Satellite bias (marine)	Global	1
Satellite bias (land)	Global + 2.5 degree SPDE	1 + 40,962
Satellite bias (ice)	Hemispheric + 2.5 degree SPDE*	2 + 40,962

ITERATIVE SOLUTION



Model Factorisation

The model is factorised into $m = 1, \dots, 3$ components that are estimated iteratively, substituting \tilde{y}_m for y :

$$\tilde{y}_m = y - \sum_{n \neq m} J_n \mu_{x_n} \tilde{y}_n$$

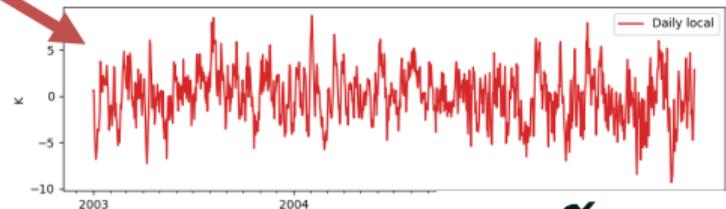
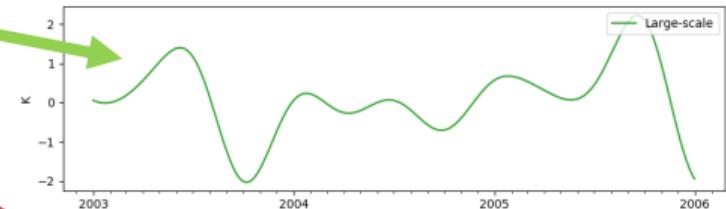
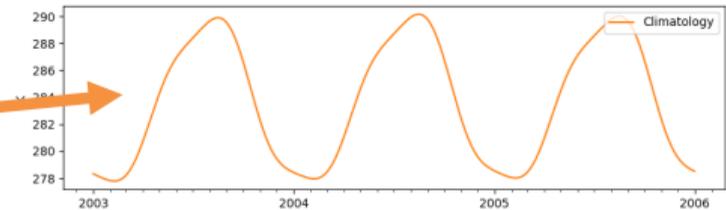
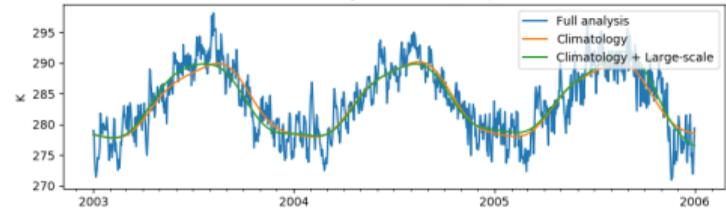
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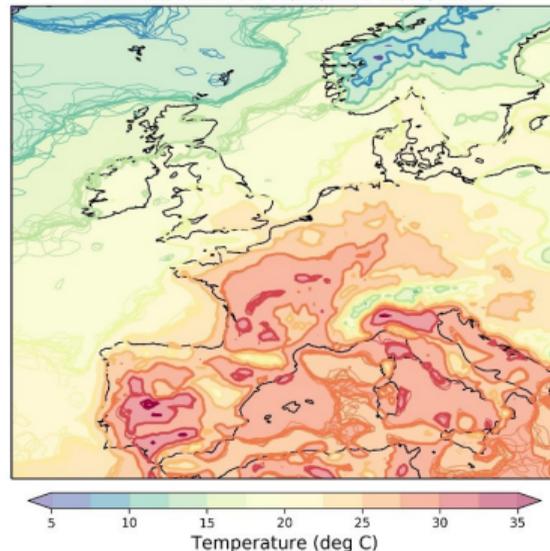
- e.g. satellite biases, station homogenisation.



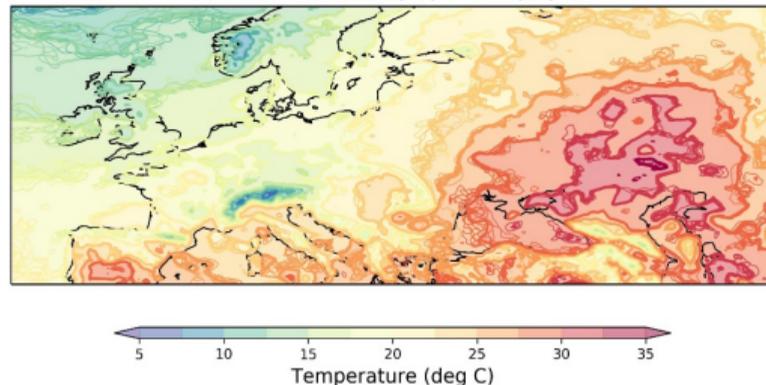
ENSEMBLE ANALYSIS

- Samples drawn from joint posterior distribution of temperature and bias variables.
- Temperature model samples projected onto analysis grid.
- Spatial/temporal correlation in analysis errors is encoded into the ensemble.
- Summary statistics can be derived from the ensemble. Expected value, total uncertainty and observation constraint information also available.

EUSTACE Ensemble 04/08/2003-13/08/2003



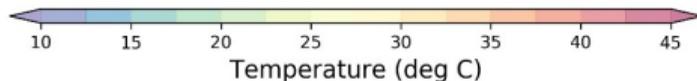
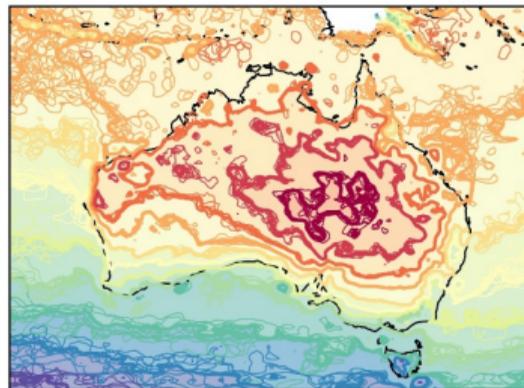
EUSTACE Ensemble 30/07/2010-05/08/2010



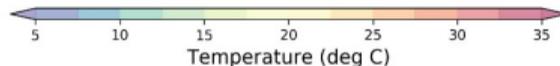
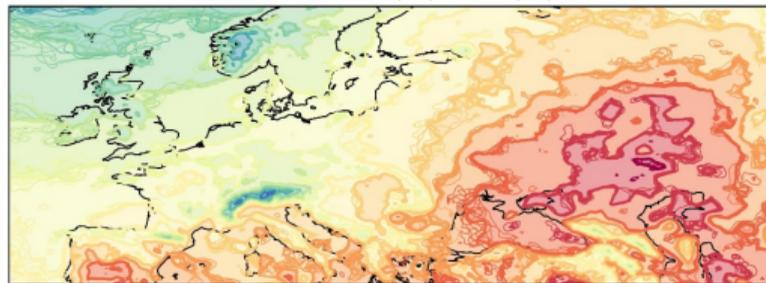
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EUSTACE Ensemble 01/01/2006-14/01/2006



EUSTACE Ensemble 30/07/2010-05/08/2010



Summary

- ▶ Challenging statistical problem, in both size and complexity
- ▶ Methods related to but different from traditional PDE solvers
- ▶ Approximate calculation techniques allows some of the complexity to be handled with reasonable computational resources
 - ▶ SPDEs and Gaussian Markov random fields
 - ▶ Fast local sparse solves
 - ▶ Global multiscale block iteration
- ▶ Close collaboration between climate scientistis, statisticians, and software engineers is essential
- ▶ Project information and links to the CEDA archive: <https://www.eustaceproject.org/>

Only partially mentioned in this talk:

- ▶ Pure conditional block updates risk getting stuck; need for convergence acceleration
- ▶ Overlapping space-time blocks for preconditioning
- ▶ Non-stationary random field parameter estimation
- ▶ Direct&iterative variance calculations to eliminate or reduce Monte Carlo error in the reconstruction uncertainties
- ▶ Fast approximate handling of correlated observation components

