Mathematics for Informatics 4a

José Figueroa-O'Farrill



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The set of outcomes is denoted Ω and the possible events define a σ -field \mathcal{F} of subsets of Ω ; that is, a family of subsets which contains \emptyset and Ω and is closed under complementation, countable union and countable intersection.

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We let N(A) denote the number of times that the event A occurs. Clearly, $0 \le N(A) \le N$. If the following limit exists

$$\lim_{N\to\infty}\frac{N(A)}{N}=\mathbb{P}(A) ,$$

the number $\mathbb{P}(A)$ obeys $0 \leq \mathbb{P}(A) \leq 1$ and is called the **probability** of *A* occuring.

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the number $\mathbb{P}(A)$ obeys $0 \leq \mathbb{P}(A) \leq 1$ and is called the **probability** of A occuring. Since $N(\Omega) = N$, we have $\mathbb{P}(\Omega) = 1$. If $A \cap B = \emptyset$, $N(A \cup B) = N(A) + N(B)$, whence $\mathbb{P}(A \cup B) = \mathbb{P}(A) + \mathbb{P}(B)$. By induction, if A_i is a finite family of pairwise disjoint events,

 $\mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_n) = \mathbb{P}(A_1) + \mathbb{P}(A_2) + \dots + \mathbb{P}(A_n) .$

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As usual, it is (theoretically) convenient to extend this to countable families.

Probability measures

Definition

A probability measure on (Ω, \mathcal{F}) is a function $\mathbb{P} : \mathcal{F} \to [0, 1]$ satisfying

• $\mathbb{P}(\Omega) = 1$

• if $A_i \in \mathcal{F}$, i = 1, 2, ... are such that $A_i \cap A_j = \emptyset$ for all $i \neq j$, then

$$\mathbb{P}\left(\bigcup_{i=1}^{\infty}A_i\right) = \sum_{i=1}^{\infty}\mathbb{P}(A_i) \ .$$

The triple $(\Omega, \mathcal{F}, \mathbb{P})$ is called a **probability space**.

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Remark

Since $\Omega = A \cup A^c$ is a disjoint union, $\mathbb{P}(A) + \mathbb{P}(A^c) = 1$. In particular, $\mathbb{P}(\emptyset) = 0$.

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Example

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Tossing a coin is a Bernoulli trial: $\Omega = \{H, T\}$.

Let us call the two outcomes generically "success" (S) and "failure" (F), so that $\Omega = \{S, F\}$. Then we have

 $\mathbb{P}(\{S\}) = p \qquad \text{and} \qquad \mathbb{P}(\{F\}) = q \;.$

Since $\Omega = \{S\} \cup \{F\}$ is a disjoint union, it follows that q = 1 - p.

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Notation

We will often drop the $\{\ \}$ when talking about events consisting of a single outcome and will write $\mathbb{P}(S) = p$ and $\mathbb{P}(F) = 1 - p$.

Fair coins and fair dice

Tossing a coin has $\Omega = \{H, T\}$. Let $\mathbb{P}(H) = p$ and $\mathbb{P}(T) = 1 - p$. The coin is fair if $p = \frac{1}{2}$, so that both H and T are equally probable.

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Similarly, a **fair** die is one where every outcome has the same probability. Since there are six outcomes, each one has probability $\frac{1}{6}$:

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Fair coins and fair dice are examples of "uniform probability spaces": those whose outcomes are all equally likely.

Uniform probability measures

Suppose that Ω is a finite set of cardinality $|\Omega|$, and suppose that every outcome is equally likely: $\mathbb{P}(\omega) = p$ for all $\omega \in \Omega$.

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Uniform probability measures

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$$1 = \mathbb{P}\left(\bigcup_{\omega \in \Omega} \omega\right) = \sum_{\omega \in \Omega} \mathbb{P}(\omega) = \sum_{\omega \in \Omega} p = p |\Omega| \text{ ,}$$

whence $p = 1/|\Omega|$.

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whence $p = 1/|\Omega|$. Now let $A \subseteq \Omega$ be an event:

$$\mathbb{P}(A) = \mathbb{P}\left(\bigcup_{\omega \in A} \omega\right) = \sum_{\omega \in A} \mathbb{P}(\omega) = \sum_{\omega \in A} \frac{1}{|\Omega|} = \frac{|A|}{|\Omega|}$$

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You draw a number at random from $\{1, 2, ..., 30\}$. What is the probability of the following events:

- A = the number drawn is even
- **2** B = the number drawn is divisible by 3
- C = the number drawn is less than 12

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$$A = \{2, 4, 6, \dots, 30\}, \text{ so } |A| = 15 \text{ and hence } \mathbb{P}(A) = \frac{15}{30} = \frac{1}{2}.$$

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1 $A = \{2, 4, 6, \dots, 30\}$, so |A| = 15 and hence $\mathbb{P}(A) = \frac{15}{30} = \frac{1}{2}$. **2** $B = \{3, 6, 9, \dots, 30\}$, so |B| = 10 and hence $\mathbb{P}(B) = \frac{10}{30} = \frac{1}{3}$.

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, so $|A| = 15$ and hence $\mathbb{P}(A) = \frac{15}{30} = \frac{1}{2}$.
• $B = \{3, 6, 9, ..., 30\}$, so $|B| = 10$ and hence $\mathbb{P}(B) = \frac{10}{30} = \frac{1}{3}$.
• $C = \{1, 2, 3, ..., 11\}$, so $|C| = 11$ and hence $\mathbb{P}(C) = \frac{11}{30}$.

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Three fair dice are rolled and their scores added. Which is more likely: a 9 or a 10?

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We could have answered this without enumeration. The average score of rolling three dice is $10\frac{1}{2}$. Since 10 is closer than 9 to the average, $\mathbb{P}(10) \ge \mathbb{P}(9)$ as a consequence of the "central limit theorem".

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Example (Alice and Bob's game)

Alice and Bob toss a fair coin in turn and the winner is the first one to get H. Suppose that Alice goes first and consider the three events:

- A = Alice wins
- B = Bob wins
- C = nobody wins

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Let ω_i be the outcome $\underbrace{TT \cdots T}_{i-1}$ H. Then $A = \{\omega_1, \omega_3, \omega_5, \dots\}$ and $B = \{\omega_2, \omega_4, \omega_6, \dots\}$.

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and $B = \{\omega_2, \omega_4, \omega_6, ...\}$. There is a further possible outcome ω_{∞} , corresponding to the unending game TTT ... in which nobody wins. Hence $C = \{\omega_{\infty}\}$. The (countably infinite) sample space is $\Omega = \{\omega_1, \omega_2, ..., \omega_{\infty}\}$, which is the disjoint union $\Omega = A \cup B \cup C$. Therefore $\mathbb{P}(A) + \mathbb{P}(B) + \mathbb{P}(C) = 1$.

There are 2^n possible outcomes of tossing the coin n times, all equally likely. Hence $\mathbb{P}(\omega_n) = 1/2^n$.

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Similarly, since $B = \bigcup_{n=1}^{\infty} \{\omega_{2n}\}$ is also a disjoint union,

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Finally, since $\mathbb{P}(A) + \mathbb{P}(B) = 1$, we see that $\mathbb{P}(C) = 0$.

Warning

Although $\mathbb{P}(\mathbb{C}) = 0$, the event \mathbb{C} is **not** impossible.

Basic properties of probability measures

Theorem

2 if $B \supseteq A$ then $\mathbb{P}(B) \ge \mathbb{P}(A)$

Basic properties of probability measures

Theorem

$$\mathbb{D} \mathbb{P}(A^c) = \mathbf{1} - \mathbb{P}(A)$$

2 if $B \supseteq A$ then $\mathbb{P}(B) \ge \mathbb{P}(A)$

Proof.

() $\Omega = A \cup A^c$ is a disjoint union, whence

 $1 = \mathbb{P}(\Omega) = \mathbb{P}(A) + \mathbb{P}(A^c)$.

Basic properties of probability measures

Theorem 2 if $B \supset A$ then $\mathbb{P}(B) \ge \mathbb{P}(A)$ Proof. **1** $\Omega = A \cup A^c$ is a disjoint union, whence $1 = \mathbb{P}(\Omega) = \mathbb{P}(A) + \mathbb{P}(A^c)$.

2 Write B as the disjoint union $B = (B \setminus A) \cup A$, whence

 $\mathbb{P}(B) = \mathbb{P}(B \setminus A) + \mathbb{P}(A) \geqslant \mathbb{P}(A)$.

Example (The Birthday problem)

What is the probability that among n people chosen at random, there are at least 2 people sharing the same birthday?

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Let A_n be the event where at least two people in n share the same birthday. Then A_n^c is the event that no two people in n share the same birthday and $\mathbb{P}(A_n) = 1 - \mathbb{P}(A_n^c)$.

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Let A_n be the event where at least two people in n share the same birthday. Then A_n^c is the event that no two people in n share the same birthday and $\mathbb{P}(A_n) = 1 - \mathbb{P}(A_n^c)$. There are 365^n possible outcomes to the birthdays of n people and $365 \times 364 \times \cdots \times (365 - n + 1)$ possible outcomes consisting of n different birthdays, hence

$$\mathbb{P}(A_n^c) = \frac{365 \times 364 \times \dots \times (365 - n + 1)}{365^n} = \prod_{i=1}^{n-1} \left(1 - \frac{i}{365}\right)$$

and

$$\mathbb{P}(A_n) = 1 - \prod_{i=1}^{n-1} \left(1 - \frac{i}{365} \right) \; .$$

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For which value of n is the chance of two people sharing the same birthday better than evens?

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• $A \cup B = (A \triangle B) \cup (A \cap B)$ whence

 $\mathbb{P}(A \cup B) = \mathbb{P}(A \triangle B) + \mathbb{P}(A \cap B)$ = $\mathbb{P}(A \setminus B) + \mathbb{P}(B \setminus A) + \mathbb{P}(A \cap B)$ = $\mathbb{P}(A) - \mathbb{P}(A \cap B) + \mathbb{P}(B) - \mathbb{P}(A \cap B) + \mathbb{P}(A \cap B)$ = $\mathbb{P}(A) + \mathbb{P}(B) - \mathbb{P}(A \cap B)$

Historical meteorological records for a certain seaside location show that on New Year's day there is a 30% chance of rain, 40% chance of being windy and 20% chance of both rain and wind. What is the chance of it being dry? dry and windy? wet or windy?

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• $\mathbb{P}(dry) = 1 - \mathbb{P}(wet) = 1 - \frac{3}{10} = \frac{7}{10}$

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- $\mathbb{P}(\text{wet or windy}) = \mathbb{P}(\text{wet}) + \mathbb{P}(\text{windy}) \mathbb{P}(\text{wet and windy}) = \frac{3}{10} + \frac{4}{10} \frac{2}{10} = \frac{1}{2}$

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Boole's inequality

Theorem

 $\mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_n) \leqslant \mathbb{P}(A_1) + \mathbb{P}(A_2) + \dots + \mathbb{P}(A_n)$

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Boole's inequality

Theorem

$$\mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_n) \leqslant \mathbb{P}(A_1) + \mathbb{P}(A_2) + \dots + \mathbb{P}(A_n)$$

Proof.

 $A_1 \cup A_2 \cup \cdots \cup A_n = (A_1 \cup A_2 \cup \cdots \cup A_{n-1}) \cup A_n$, and by the inclusion-exclusion rule,

$$\mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_n) \leqslant \mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_{n-1}) + \mathbb{P}(A_n)$$

But now $A_1 \cup A_2 \cup \cdots \cup A_{n-1} = (A_1 \cup A_2 \cup \cdots \cup A_{n-2}) \cup A_{n-1}$, so that

 $\mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_n) \leqslant \mathbb{P}(A_1 \cup A_2 \cup \dots \cup A_{n-2}) + \mathbb{P}(A_{n-1}) + \mathbb{P}(A_n)$

et cetera.

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General inclusion-exclusion rule

Theorem $\mathbb{P}\left(\bigcup_{i=1}^{n} A_{i}\right) = \sum_{i=1}^{n} \mathbb{P}(A_{i}) - \sum_{1 \leq i < j \geq n} \mathbb{P}(A_{i} \cap A_{j})$ $+ \sum_{1 \leq i < j < k \leq n} \mathbb{P}(A_{i} \cap A_{j} \cap A_{k}) - \dots + (-1)^{n} \mathbb{P}\left(\bigcap_{i=1}^{n} A_{i}\right)$

Proof.

Use induction from the simple inclusion-exclusion rule.

Continuity

Theorem

• Let
$$A_1 \subseteq A_2 \subseteq A_3 \subseteq \cdots$$
 and let $A = \bigcup_{i=1}^{\infty} A_i = \lim_{i \to \infty} A_i$.
Then $P(A) = \lim_{i \to \infty} \mathbb{P}(A_i)$.

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- 2 Let $B_1 \supseteq B_2 \supseteq B_3 \supseteq \cdots$ and let $B = \bigcap_{i=1}^{\infty} B_i = \lim_{i \to \infty} B_i$. Then $P(B) = \lim_{i \to \infty} \mathbb{P}(B_i)$.

Proof.

$$\begin{split} \mathbb{P}(A) &= \mathbb{P}(A_1) + \mathbb{P}(A_2 \setminus A_1) + \mathbb{P}(A_3 \setminus A_2) + \cdots \\ &= \mathbb{P}(A_1) + (\mathbb{P}(A_2) - \mathbb{P}(A_1)) + (\mathbb{P}(A_3) - \mathbb{P}(A_2)) + \cdots \\ &= \lim_{n \to \infty} \mathbb{P}(A_n). \end{split}$$

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Take complements of the previous proof.

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Probability spaces with Ω finite and \mathbb{P} uniformly distributed ("all outcomes equally likely") are particularly amenable to counting techniques from combinatorial analysis.

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