Mathematics for Informatics 4a

José Figueroa-O'Farrill



Lecture 3 25 January 2012

< ∃ >

With every experiment we associate a probability space $(\Omega, \mathcal{F}, \mathbb{P})$, where

With every experiment we associate a probability space $(\Omega, \mathcal{F}, \mathbb{P})$, where

• Ω is the set of all possible outcomes of the experiment;

With every experiment we associate a probability space $(\Omega, \mathcal{F}, \mathbb{P})$, where

- Ω is the set of all possible outcomes of the experiment;
- F is a σ-field of subsets of Ω: containing Ω and closed under complementation and countable unions; and

くぼう くほう くほう

With every experiment we associate a probability space $(\Omega, \mathcal{F}, \mathbb{P})$, where

- Ω is the set of all possible outcomes of the experiment;
- F is a σ-field of subsets of Ω: containing Ω and closed under complementation and countable unions; and
- 𝔅 𝔅 → [0, 1] is a function normalised to ℙ(Ω) = 1 and countably additive over disjoint unions.

э.

く 同 と く ヨ と く ヨ と

With every experiment we associate a probability space $(\Omega, \mathcal{F}, \mathbb{P})$, where

- Ω is the set of all possible outcomes of the experiment;
- F is a σ-field of subsets of Ω: containing Ω and closed under complementation and countable unions; and
- 𝔅 𝔅 → [0, 1] is a function normalised to ℙ(Ω) = 1 and countably additive over disjoint unions.

We also introduced uniform probability spaces with Ω a finite set and $\mathbb{P}(A) = |A|/|\Omega|$ for every event *A*.

▲□ ▶ ▲ □ ▶ ▲ □ ▶ □ ● ● ● ●

With every experiment we associate a probability space $(\Omega, \mathcal{F}, \mathbb{P})$, where

- Ω is the set of all possible outcomes of the experiment;
- F is a σ-field of subsets of Ω: containing Ω and closed under complementation and countable unions; and
- 𝔅 𝔅 → [0, 1] is a function normalised to ℙ(Ω) = 1 and countably additive over disjoint unions.

We also introduced uniform probability spaces with Ω a finite set and $\mathbb{P}(A) = |A|/|\Omega|$ for every event *A*.

The thought of the day

Probability is a measure of our ignorance and hence, when our knowledge about a system changes, the probability of events should also change to reflect the new knowledge.

э.

ヘロト 人間 とくほ とくほ とう

We roll two fair dice in turn. Let A be the event "the second die shows a higher score than the first". What is $\mathbb{P}(A)$?

・ロン ・日 ・ ・ ヨン ・ ヨン

We roll two fair dice in turn. Let A be the event "the second die shows a higher score than the first". What is $\mathbb{P}(A)$? There are 6² possible outcomes of which $\binom{6}{2} = 15$ lie in A. Hence $\mathbb{P}(A) = \frac{15}{36} = \frac{5}{12}$.

イロト 不得 トイヨト イヨト 二日

We roll two fair dice in turn. Let A be the event "the second die shows a higher score than the first". What is $\mathbb{P}(A)$? There are 6^2 possible outcomes of which $\binom{6}{2} = 15$ lie in A. Hence $\mathbb{P}(A) = \frac{15}{36} = \frac{5}{12}$.

Now suppose that the first die turns out to be \mathfrak{U} . What is $\mathbb{P}(A)$?

We roll two fair dice in turn. Let A be the event "the second die shows a higher score than the first". What is $\mathbb{P}(A)$? There are 6^2 possible outcomes of which $\binom{6}{2} = 15$ lie in A. Hence $\mathbb{P}(A) = \frac{15}{36} = \frac{5}{12}$. Now suppose that the first die turns out to be I. What is $\mathbb{P}(A)$?

There is now only one positive outcome: namely, (:, :).

We roll two fair dice in turn. Let A be the event "the second die shows a higher score than the first". What is $\mathbb{P}(A)$? There are 6² possible outcomes of which $\binom{6}{2} = 15$ lie in A. Hence $\mathbb{P}(A) = \frac{15}{36} = \frac{5}{12}$. Now suppose that the first die turns out to be I. What is $\mathbb{P}(A)$? There is now only one positive outcome: namely, $(\textcircled{I}, \fbox{I})$. But the sample space has changed as well. Once we know that the first die is I, the sample space consists

of six outcomes:

 $\{(\overleftarrow{\bullet}, \bullet), \ (\overleftarrow{\bullet}, \bullet)\}$

and hence $\mathbb{P}(A) = \frac{1}{6}$.

Alice and Bob have two children. What is the probability that they are both girls?

Alice and Bob have two children. What is the probability that they are both girls?

Assuming a uniform distribution (i.e., boy or girl is equally likely), every outcome in our sample space

 $\Omega = \{(\texttt{q},\texttt{q}),(\texttt{q},\texttt{d}'),(\texttt{d}',\texttt{q}),(\texttt{d}',\texttt{d}')\}$

has probability $\frac{1}{4}$.

・ 同 ト ・ ヨ ト ・ ヨ ト

3

4/23

Alice and Bob have two children. What is the probability that they are both girls?

Assuming a uniform distribution (i.e., boy or girl is equally likely), every outcome in our sample space

 $\Omega = \{(\texttt{q},\texttt{q}), (\texttt{q},\texttt{d}), (\texttt{d},\texttt{q}), (\texttt{d},\texttt{d})\}$

has probability $\frac{1}{4}$. The desired outcome A = (Q, Q) has $\mathbb{P}(A) = \frac{1}{4}$.

・ 同 ト ・ ヨ ト ・ ヨ ト

Alice and Bob have two children. What is the probability that they are both girls?

Assuming a uniform distribution (i.e., boy or girl is equally likely), every outcome in our sample space

 $\Omega = \{(\texttt{q},\texttt{q}),(\texttt{q},\texttt{d}^{\texttt{r}}),(\texttt{d}^{\texttt{r}},\texttt{q}),(\texttt{d}^{\texttt{r}},\texttt{d}^{\texttt{r}})\}$

has probability $\frac{1}{4}$. The desired outcome $A = (\varphi, \varphi)$ has $\mathbb{P}(A) = \frac{1}{4}$. Now suppose that we know that one of the children is a girl. What is the probability now?

くぼう くほう くほう

Alice and Bob have two children. What is the probability that they are both girls?

Assuming a uniform distribution (i.e., boy or girl is equally likely), every outcome in our sample space

 $\Omega = \{(\texttt{q},\texttt{q}),(\texttt{q},\texttt{d}^{\texttt{r}}),(\texttt{d}^{\texttt{r}},\texttt{q}),(\texttt{d}^{\texttt{r}},\texttt{d}^{\texttt{r}})\}$

has probability $\frac{1}{4}$. The desired outcome $A = (\varphi, \varphi)$ has $\mathbb{P}(A) = \frac{1}{4}$. Now suppose that we know that one of the children is a girl. What is the probability now?

The sample space is now $\{(\varphi, \varphi), (\varphi, \sigma^2), (\sigma^2, \varphi)\}$, whereas there is still only one positive outcome, whence $\mathbb{P}(A) = \frac{1}{3}$.

・ロン ・日 ・ ・ ヨン ・ ヨン

Alice and Bob have two children. What is the probability that they are both girls?

Assuming a uniform distribution (i.e., boy or girl is equally likely), every outcome in our sample space

 $\Omega = \{(\texttt{q},\texttt{q}),(\texttt{q},\texttt{d}^{\texttt{r}}),(\texttt{d}^{\texttt{r}},\texttt{q}),(\texttt{d}^{\texttt{r}},\texttt{d}^{\texttt{r}})\}$

has probability $\frac{1}{4}$. The desired outcome $A = (\varphi, \varphi)$ has $\mathbb{P}(A) = \frac{1}{4}$. Now suppose that we know that one of the children is a girl. What is the probability now?

The sample space is now $\{(\varphi, \varphi), (\varphi, \sigma^2), (\sigma^2, \varphi)\}$, whereas there is still only one positive outcome, whence $\mathbb{P}(A) = \frac{1}{3}$.

What about if we know that the oldest child is a girl?

▲御 → ▲ 唐 → ▲ 唐 → □

Alice and Bob have two children. What is the probability that they are both girls?

Assuming a uniform distribution (i.e., boy or girl is equally likely), every outcome in our sample space

 $\Omega = \{(\texttt{q},\texttt{q}),(\texttt{q},\texttt{d}^{\texttt{r}}),(\texttt{d}^{\texttt{r}},\texttt{q}),(\texttt{d}^{\texttt{r}},\texttt{d}^{\texttt{r}})\}$

has probability $\frac{1}{4}$. The desired outcome $A = (\varphi, \varphi)$ has $\mathbb{P}(A) = \frac{1}{4}$. Now suppose that we know that one of the children is a girl. What is the probability now?

The sample space is now $\{(\varphi, \varphi), (\varphi, \sigma^2), (\sigma^2, \varphi)\}$, whereas there is still only one positive outcome, whence $\mathbb{P}(A) = \frac{1}{3}$.

What about if we know that the oldest child is a girl?

The sample space is now $\{(\varphi, \varphi), (\varphi, \sigma^2)\}$, whence $\mathbb{P}(A) = \frac{1}{2}$.

イロト 不得 トイヨト イヨト

Conditional probability

Definition

Let A, B be events. The **conditional probability** $\mathbb{P}(A|B)$ of the event of "A occurs given that B occured" is

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$$

assuming that $\mathbb{P}(B) > 0$.

Conditional probability

Definition

Let A, B be events. The **conditional probability** $\mathbb{P}(A|B)$ of the event of "A occurs given that B occured" is

 $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$

assuming that $\mathbb{P}(B) > 0$.

Warning!

 $\mathbb{P}(A|B) \neq \mathbb{P}(B|A)$, unless $\mathbb{P}(A) = \mathbb{P}(B)$. The incorrect equality formed the basis of the prosecution's case in the infamous *affaire Dreyfus* in France at the turn of the 20th century. To find out more, visit my colleague Andrew Ranicki's page on this subject: http://www.maths.ed.ac.uk/~aar/dreyfus.htm.

3

・ロト ・四ト ・ヨト ・ ヨト



This Dreyfus...

2

・ロト ・ 四ト ・ ヨト ・ ヨト



・ロト ・ 四ト ・ ヨト ・ ヨト

... not this Dreyfus!

José Figueroa-O'Farrill mi4a (Probability) Lecture 3

Э.

Conditional probability from uniform probability

If $(\Omega, \mathcal{F}, \mathbb{P})$ is a uniform probability space (e.g., Ω finite and all outcomes equally probable) and A, B are two events, we have

$$\mathbb{P}(A) = \frac{|A|}{|\Omega|}$$
 and $\mathbb{P}(B) = \frac{|B|}{|\Omega|}$.

・ 同 ト ・ ヨ ト ・ ヨ ト …

Conditional probability from uniform probability

If $(\Omega, \mathcal{F}, \mathbb{P})$ is a uniform probability space (e.g., Ω finite and all outcomes equally probable) and A, B are two events, we have

$$\mathbb{P}(A) = \frac{|A|}{|\Omega|}$$
 and $\mathbb{P}(B) = \frac{|B|}{|\Omega|}$.

If B occurs, the possible outcomes are those in B and they remain equally likely with probability 1/|B|.

3

・ 同 ト ・ ヨ ト ・ ヨ ト …

Conditional probability from uniform probability

If $(\Omega, \mathcal{F}, \mathbb{P})$ is a uniform probability space (e.g., Ω finite and all outcomes equally probable) and A, B are two events, we have

$$\mathbb{P}(A) = \frac{|A|}{|\Omega|}$$
 and $\mathbb{P}(B) = \frac{|B|}{|\Omega|}$.

If B occurs, the possible outcomes are those in B and they remain equally likely with probability 1/|B|.

The event A occurs if and only if $A \cap B$ occurs, whence the probability that A occurs given that B occurs is

$$\frac{|A \cap B|}{|B|} = \frac{|A \cap B|}{|\Omega|} / \frac{|B|}{|\Omega|} = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \mathbb{P}(A|B) .$$

3

・ 同 ト ・ ヨ ト ・ ヨ ト

We repeat the experiment N times, with B occuring N(B) times.

We repeat the experiment N times, with B occuring N(B) times. Since we know that B occurs we are in one of those N(B) trials.

くぼう くほう くほう

8/23

We repeat the experiment N times, with B occuring N(B) times. Since we know that B occurs we are in one of those N(B) trials. The event A occurs in N(A \cap B) of them, so the probability of A occuring given that B does is given by the limit N $\rightarrow \infty$ of

$$\frac{\mathsf{N}(A \cap B)}{\mathsf{N}(B)} = \frac{\mathsf{N}(A \cap B)}{\mathsf{N}} \Big/ \frac{\mathsf{N}(B)}{\mathsf{N}} \to \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \mathbb{P}(A|B) \; .$$

э.

く 同 と く ヨ と く ヨ と

We repeat the experiment N times, with B occuring N(B) times. Since we know that B occurs we are in one of those N(B) trials. The event A occurs in N(A \cap B) of them, so the probability of A occuring given that B does is given by the limit N $\rightarrow \infty$ of

$$\frac{\mathsf{N}(\mathsf{A}\cap\mathsf{B})}{\mathsf{N}(\mathsf{B})} = \frac{\mathsf{N}(\mathsf{A}\cap\mathsf{B})}{\mathsf{N}} \bigg/ \frac{\mathsf{N}(\mathsf{B})}{\mathsf{N}} \to \frac{\mathbb{P}(\mathsf{A}\cap\mathsf{B})}{\mathbb{P}(\mathsf{B})} = \mathbb{P}(\mathsf{A}|\mathsf{B}) \; .$$

We will now revisit the previous examples in the language of conditional probability.

くロン 不得 とくほう くほう 二日 二

Let *A* be the event "the second die shows a greater score than the first" and let *B* be the event "the first die shows a ".

Let A be the event "the second die shows a greater score than the first" and let B be the event "the first die shows a ". Then

 $B = \{ (\textcircled{i}, \textcircled{i}), (\textcircled{i}, \textcircled{i}), (\textcircled{i}, \textcircled{i}), (\textcircled{i}, \textcircled{i}), (\textcircled{i}, \textcircled{i}), (\textcircled{i}, \textcircled{i}) \}$ $A \cap B = \{ (\textcircled{i}, \textcircled{i}) \}$

whence

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \frac{|A \cap B|}{|B|} = \frac{1}{6}.$$

э.

Let

$$\begin{split} A &= \{(\wp, \wp)\}\\ B &= \{(\wp, \wp), (\wp, \sigma^*), (\sigma^*, \wp)\}\\ C &= \{(\wp, \wp), (\wp, \sigma^*)\} \end{split}$$

"both are girls" "at least one is a girl" "the oldest is a girl" .

Let

$$\begin{split} A &= \{(\wp, \wp)\}\\ B &= \{(\wp, \wp), (\wp, \sigma^*), (\sigma^*, \wp)\}\\ C &= \{(\wp, \wp), (\wp, \sigma^*)\} \end{split}$$

"both are girls" "at least one is a girl" "the oldest is a girl" .

Then

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \frac{|A \cap B|}{|B|} = \frac{1}{3}$$

Let

$$A = \{(\varphi, \varphi)\}$$
$$B = \{(\varphi, \varphi), (\varphi, \sigma^2), (\sigma^2, \varphi)\}$$
$$C = \{(\varphi, \varphi), (\varphi, \sigma^2)\}$$

"both are girls" "at least one is a girl" "the oldest is a girl" .

Then

and

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)} = \frac{|A \cap B|}{|B|} = \frac{1}{3}$$
$$\mathbb{P}(A|C) = \frac{\mathbb{P}(A \cap C)}{\mathbb{P}(C)} = \frac{|A \cap C|}{|C|} = \frac{1}{2}$$

Example

A box contains a double-headed coin, a double-tailed coin and a conventional coin. A coin is picked at random, tossed and the result is a head. What is the probability that it is the double-headed coin?

・ 同 ト ・ ヨ ト ・ ヨ ト ・
A box contains a double-headed coin, a double-tailed coin and a conventional coin. A coin is picked at random, tossed and the result is a head. What is the probability that it is the double-headed coin?

Let D be the event that we did pick the double-headed coin and let H be the event that the coin we picked and tossed, came up heads. We want to calculate $\mathbb{P}(D|H)$.

A D A D A D A

A box contains a double-headed coin, a double-tailed coin and a conventional coin. A coin is picked at random, tossed and the result is a head. What is the probability that it is the double-headed coin?

Let D be the event that we did pick the double-headed coin and let H be the event that the coin we picked and tossed, came up heads. We want to calculate $\mathbb{P}(D|H)$.

There are 3 coins and hence 6 possible outcomes, of which 3 are heads. Therefore $\mathbb{P}(H)=\frac{3}{6}=\frac{1}{2}.$

伺下 イヨト イヨト

A box contains a double-headed coin, a double-tailed coin and a conventional coin. A coin is picked at random, tossed and the result is a head. What is the probability that it is the double-headed coin?

Let D be the event that we did pick the double-headed coin and let H be the event that the coin we picked and tossed, came up heads. We want to calculate $\mathbb{P}(D|H)$.

There are 3 coins and hence 6 possible outcomes, of which 3 are heads. Therefore $\mathbb{P}(H) = \frac{3}{6} = \frac{1}{2}$.

Of those 3 outcomes, two come from picking the double-headed coin, whence $\mathbb{P}(D \cap H) = \frac{2}{6} = \frac{1}{3}$.

< 回 > < 回 > < 回 > -

A box contains a double-headed coin, a double-tailed coin and a conventional coin. A coin is picked at random, tossed and the result is a head. What is the probability that it is the double-headed coin?

Let D be the event that we did pick the double-headed coin and let H be the event that the coin we picked and tossed, came up heads. We want to calculate $\mathbb{P}(D|H)$.

There are 3 coins and hence 6 possible outcomes, of which 3 are heads. Therefore $\mathbb{P}(H) = \frac{3}{6} = \frac{1}{2}$.

Of those 3 outcomes, two come from picking the double-headed coin, whence $\mathbb{P}(D \cap H) = \frac{2}{6} = \frac{1}{3}$. Therefore

$$\mathbb{P}(\mathsf{D}|\mathsf{H}) = \frac{\mathbb{P}(\mathsf{D}\cap\mathsf{H})}{\mathbb{P}(\mathsf{H})} = \frac{1}{3} / \frac{1}{2} = \frac{2}{3} \; .$$

< 回 > < 回 > < 回 > -

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$.

イロト イヨト イヨト -

э.

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$. Let $\mathbb{P}'(A) := \mathbb{P}(A|B)$.

くロン (雪) (ヨ) (ヨ)

э.

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$. Let $\mathbb{P}'(A) := \mathbb{P}(A|B)$.

Then \mathbb{P}' is again a probability measure:

• $\mathbb{P}': \mathfrak{F} \to [0, 1]$

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$. Let $\mathbb{P}'(A) := \mathbb{P}(A|B)$.

Then \mathbb{P}' is again a probability measure:

• $\mathbb{P}': \mathcal{F} \to [0, 1]$

• $\mathbb{P}'(\Omega) = \mathbf{1}$

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$. Let $\mathbb{P}'(A) := \mathbb{P}(A|B)$.

Then \mathbb{P}' is again a probability measure:

- $\mathbb{P}': \mathcal{F} \to [0, 1]$
- $\mathbb{P}'(\Omega) = \mathbf{1}$

ヘロン 人間 とくほ とくほ とう

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$. Let $\mathbb{P}'(A) := \mathbb{P}(A|B)$.

Then \mathbb{P}' is again a probability measure:

- $\mathbb{P}': \mathcal{F} \to [0, 1]$
- $\mathbb{P}'(\Omega) = \mathbf{1}$

This means that all results proved for general probability measures apply to the conditional probability as well.

ヘロト 人間 とくほ とくほ とう

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$. Let $\mathbb{P}'(A) := \mathbb{P}(A|B)$.

Then \mathbb{P}' is again a probability measure:

- $\mathbb{P}': \mathcal{F} \to [0, 1]$
- $\mathbb{P}'(\Omega) = 1$
- P' is countably additive over disjoint unions

This means that all results proved for general probability measures apply to the conditional probability as well. In fact, we can define a new probability space $(B, \mathcal{F}', \mathbb{P}')$, where

 $\mathfrak{F}' = \{A \cap B | A \in \mathfrak{F}\}$.

Let $(\Omega, \mathcal{F}, \mathbb{P})$ be a probability space and let $B \in \mathcal{F}$ be an event with $\mathbb{P}(B) > 0$. Let $\mathbb{P}'(A) := \mathbb{P}(A|B)$.

Then \mathbb{P}' is again a probability measure:

- $\mathbb{P}': \mathcal{F} \to [0, 1]$
- $\mathbb{P}'(\Omega) = 1$
- P' is countably additive over disjoint unions

This means that all results proved for general probability measures apply to the conditional probability as well. In fact, we can define a new probability space $(B, \mathcal{F}', \mathbb{P}')$, where

 $\mathfrak{F}' = \{ A \cap B | A \in \mathfrak{F} \}$.

Recall that \mathfrak{F}' is again a σ -field (cf. Tutorial Sheet 1), only this time of subsets of B.

From the • definition of conditional probability there follows the

Multiplication rule

 $\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B)$

which holds even if $\mathbb{P}(B) = \mathbf{0}$.

э.

From the • definition of conditional probability there follows the

Multiplication rule

$\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B)$

which holds even if $\mathbb{P}(B) = \mathbf{0}$.

Example

I have 5 red socks and 3 black socks in a drawer. I pick two socks at random. What is the probability that I get a black pair?

From the • definition of conditional probability there follows the

Multiplication rule

$$\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B)$$

which holds even if $\mathbb{P}(B) = \mathbf{0}$.

Example

I have 5 red socks and 3 black socks in a drawer. I pick two socks at random. What is the probability that I get a black pair? Let P = "the pair is black" and B = "the first sock is black". Then clearly $P \subset B$, so $P = P \cap B$.

From the • definition of conditional probability there follows the

Multiplication rule

$\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B)$

which holds even if $\mathbb{P}(B) = \mathbf{0}$.

Example

I have 5 red socks and 3 black socks in a drawer. I pick two socks at random. What is the probability that I get a black pair? Let P = "the pair is black" and B = "the first sock is black". Then clearly $P \subset B$, so $P = P \cap B$. Also $\mathbb{P}(B) = \frac{3}{8}$ and $\mathbb{P}(P|B) = \frac{2}{7}$

From the • definition of conditional probability there follows the

Multiplication rule

$$\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B)$$

which holds even if $\mathbb{P}(B) = \mathbf{0}$.

Example

I have 5 red socks and 3 black socks in a drawer. I pick two socks at random. What is the probability that I get a black pair? Let P = "the pair is black" and B = "the first sock is black". Then clearly $P \subset B$, so $P = P \cap B$. Also $\mathbb{P}(B) = \frac{3}{8}$ and $\mathbb{P}(P|B) = \frac{2}{7}$, whence

$$\mathbb{P}(\mathsf{P}) = \mathbb{P}(\mathsf{P} \cap \mathsf{B}) = \mathbb{P}(\mathsf{P}|\mathsf{B})\mathbb{P}(\mathsf{B}) = \frac{2}{7} \times \frac{3}{8} = \frac{3}{28}.$$

Extended multiplication rule

Theorem

Let A_1, A_2, \ldots, A_n be events with $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_n) > 0$. Then

$$\mathbb{P}(A_1 \cap \dots \cap A_n) = \mathbb{P}(A_n | A_{n-1} \cap \dots \cap A_1) \times \\ \times \mathbb{P}(A_{n-1} | A_{n-2} \cap \dots \cap A_1) \cdots \mathbb{P}(A_2 | A_1) \mathbb{P}(A_1)$$

イロト 不得 トイヨト イヨト

= 990

Extended multiplication rule

Theorem

Let A_1, A_2, \ldots, A_n be events with $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_n) > 0$. Then

$$\mathbb{P}(A_1 \cap \dots \cap A_n) = \mathbb{P}(A_n | A_{n-1} \cap \dots \cap A_1) \times \\ \times \mathbb{P}(A_{n-1} | A_{n-2} \cap \dots \cap A_1) \dots \mathbb{P}(A_2 | A_1) \mathbb{P}(A_1)$$

Proof.

Just use that for k = 2, ..., n,

$$\mathbb{P}(A_k|A_{k-1}\cap\cdots\cap A_1) = \frac{\mathbb{P}(A_k\cap\cdots\cap A_1)}{\mathbb{P}(A_{k-1}\cap\cdots\cap A_1)}$$

and multiply the terms in the RHS to obtain the LHS.

A box contains 15 numbered balls: six balls have the number 1, four balls the number 2 and five balls the number 3. Suppose we draw three balls without replacement. What is the probability that we draw the balls in increasing numerical order?

向下 イヨト イヨト

A box contains 15 numbered balls: six balls have the number 1, four balls the number 2 and five balls the number 3. Suppose we draw three balls without replacement. What is the probability that we draw the balls in increasing numerical order? Let A be event "first ball has number 1", B the event "second ball has number 2" and C the event "third ball has number 3". We want to compute $\mathbb{P}(A \cap B \cap C)$.

向下 イヨト イヨト

A box contains 15 numbered balls: six balls have the number 1, four balls the number 2 and five balls the number 3. Suppose we draw three balls without replacement. What is the probability that we draw the balls in increasing numerical order? Let A be event "first ball has number 1", B the event "second ball has number 2" and C the event "third ball has number 3". We want to compute $\mathbb{P}(A \cap B \cap C)$. By the extended multiplication rule,

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(C|A \cap B)\mathbb{P}(B|A)\mathbb{P}(A)$

向下 イヨト イヨト

A box contains 15 numbered balls: six balls have the number 1, four balls the number 2 and five balls the number 3. Suppose we draw three balls without replacement. What is the probability that we draw the balls in increasing numerical order? Let A be event "first ball has number 1", B the event "second ball has number 2" and C the event "third ball has number 3". We want to compute $\mathbb{P}(A \cap B \cap C)$. By the extended multiplication rule,

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(C|A \cap B)\mathbb{P}(B|A)\mathbb{P}(A)$

where $\mathbb{P}(A) = \frac{6}{15} = \frac{2}{5}$, $\mathbb{P}(B|A) = \frac{4}{14} = \frac{2}{7}$ and $\mathbb{P}(C|A \cap B) = \frac{5}{13}$

< 回 > < 三 > < 三 > 、

A box contains 15 numbered balls: six balls have the number 1, four balls the number 2 and five balls the number 3. Suppose we draw three balls without replacement. What is the probability that we draw the balls in increasing numerical order? Let A be event "first ball has number 1", B the event "second ball has number 2" and C the event "third ball has number 3". We want to compute $\mathbb{P}(A \cap B \cap C)$. By the extended multiplication rule,

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(C|A \cap B)\mathbb{P}(B|A)\mathbb{P}(A)$

where $\mathbb{P}(A) = \frac{6}{15} = \frac{2}{5}$, $\mathbb{P}(B|A) = \frac{4}{14} = \frac{2}{7}$ and $\mathbb{P}(C|A \cap B) = \frac{5}{13}$, whence $\mathbb{P}(A \cap B \cap C) = \frac{2}{5} \times \frac{2}{7} \times \frac{5}{12} = \frac{4}{01} \simeq 4.4\%$

ヘロト 人間 とくほとく ほとう

A bag contains 26 cards, each one with a letter in $\{a, b, c, ..., z\}$. You take 7 cards at random without replacement. What is the probability that you can spell "dreyfus" with them?

A bag contains 26 cards, each one with a letter in {a, b, c, ..., z}. You take 7 cards at random without replacement. What is the probability that you can spell "dreyfus" with them? Let A_i be the event that the ith card is one with a letter in {d, r, e, y, f, u, s}. We are after $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_7)$.

・ 同 ト ・ ヨ ト ・ ヨ ト …

э.

A bag contains 26 cards, each one with a letter in {a, b, c, ..., z}. You take 7 cards at random without replacement. What is the probability that you can spell "dreyfus" with them? Let A_i be the event that the ith card is one with a letter in {d, r, e, y, f, u, s}. We are after $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_7)$. We use the extended multiplication rule:

 $\mathbb{P}(A_1 \cap \dots \cap A_7) = \mathbb{P}(A_1)\mathbb{P}(A_2|A_1) \cdots \mathbb{P}(A_7|A_6 \cap \dots \cap A_1)$

・ロン ・雪 と ・ ヨ と ・ ヨ と

A bag contains 26 cards, each one with a letter in {a, b, c, ..., z}. You take 7 cards at random without replacement. What is the probability that you can spell "dreyfus" with them? Let A_i be the event that the ith card is one with a letter in {d, r, e, y, f, u, s}. We are after $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_7)$. We use the extended multiplication rule:

 $\mathbb{P}(A_1 \cap \dots \cap A_7) = \mathbb{P}(A_1)\mathbb{P}(A_2|A_1) \cdots \mathbb{P}(A_7|A_6 \cap \dots \cap A_1)$ $= \frac{7}{26} \times \frac{6}{25} \times \frac{5}{24} \times \dots \times \frac{1}{20}$

イロト 不得 トイヨト イヨト

A bag contains 26 cards, each one with a letter in {a, b, c, ..., z}. You take 7 cards at random without replacement. What is the probability that you can spell "dreyfus" with them? Let A_i be the event that the ith card is one with a letter in {d, r, e, y, f, u, s}. We are after $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_7)$. We use the extended multiplication rule:

 $\mathbb{P}(A_1 \cap \dots \cap A_7) = \mathbb{P}(A_1)\mathbb{P}(A_2|A_1) \cdots \mathbb{P}(A_7|A_6 \cap \dots \cap A_1)$ $= \frac{7}{26} \times \frac{6}{25} \times \frac{5}{24} \times \dots \times \frac{1}{20}$ $= \frac{7 \times 6 \times \dots \times 1}{26 \times 25 \times \dots \times 20} \times \frac{19!}{19!}$

・ロン ・雪 と ・ ヨ と ・ ヨ と

A bag contains 26 cards, each one with a letter in {a, b, c, ..., z}. You take 7 cards at random without replacement. What is the probability that you can spell "dreyfus" with them? Let A_i be the event that the ith card is one with a letter in {d, r, e, y, f, u, s}. We are after $\mathbb{P}(A_1 \cap A_2 \cap \cdots \cap A_7)$. We use the extended multiplication rule:

$$\mathbb{P}(A_1 \cap \dots \cap A_7) = \mathbb{P}(A_1)\mathbb{P}(A_2|A_1) \cdots \mathbb{P}(A_7|A_6 \cap \dots \cap A_1)$$
$$= \frac{7}{26} \times \frac{6}{25} \times \frac{5}{24} \times \dots \times \frac{1}{20}$$
$$= \frac{7 \times 6 \times \dots \times 1}{26 \times 25 \times \dots \times 20} \times \frac{19!}{19!}$$
$$= \binom{26}{7}^{-1} = \frac{1}{657800}.$$

Heuristically, two events A, B are said to be independent if the chance of one occurring is not altered by the other's occurrence; that is, $\mathbb{P}(A|B) = \mathbb{P}(A)$.

くぼう くほう くほう

Heuristically, two events A, B are said to be independent if the chance of one occurring is not altered by the other's occurrence; that is, $\mathbb{P}(A|B) = \mathbb{P}(A)$. The multiplication rule the implies that

 $\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B) = \mathbb{P}(A)\mathbb{P}(B)$

and, by symmetry, $\mathbb{P}(B|A) = \mathbb{P}(B)$.

э.

Heuristically, two events A, B are said to be independent if the chance of one occurring is not altered by the other's occurrence; that is, $\mathbb{P}(A|B) = \mathbb{P}(A)$. The multiplication rule the implies that

 $\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B) = \mathbb{P}(A)\mathbb{P}(B)$

and, by symmetry, $\mathbb{P}(B|A) = \mathbb{P}(B)$.

Definition

Two events A, B are independent if

 $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$

く 伺 とう きょう く き とう

э.

Heuristically, two events A, B are said to be independent if the chance of one occurring is not altered by the other's occurrence; that is, $\mathbb{P}(A|B) = \mathbb{P}(A)$. The multiplication rule the implies that

 $\mathbb{P}(A \cap B) = \mathbb{P}(A|B)\mathbb{P}(B) = \mathbb{P}(A)\mathbb{P}(B)$

and, by symmetry, $\mathbb{P}(B|A) = \mathbb{P}(B)$.

Definition

Two events A, B are independent if

 $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$

Example

"Not winning the lottery last week" and "winning the lottery this week" are independent events.

Suppose that we have three events *A*, *B*, *C*. We say that they are **independent** if any two are independent:

 $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B) \qquad \mathbb{P}(A \cap C) = \mathbb{P}(A)\mathbb{P}(C) \qquad \mathbb{P}(B \cap C) = \mathbb{P}(B)\mathbb{P}(C)$

イロト 不得 トイヨト イヨト

э.

Suppose that we have three events *A*, *B*, *C*. We say that they are **independent** if any two are independent:

 $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B) \qquad \mathbb{P}(A \cap C) = \mathbb{P}(A)\mathbb{P}(C) \qquad \mathbb{P}(B \cap C) = \mathbb{P}(B)\mathbb{P}(C)$

and in addition

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C)$.

イロト 不得 トイヨト イヨト
Suppose that we have three events *A*, *B*, *C*. We say that they are **independent** if any two are independent:

 $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B) \qquad \mathbb{P}(A \cap C) = \mathbb{P}(A)\mathbb{P}(C) \qquad \mathbb{P}(B \cap C) = \mathbb{P}(B)\mathbb{P}(C)$

and in addition

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C)$.

Warning!

Neither of the two sets of conditions imply the other.

イロト 不得 トイヨト イヨト

Suppose that we have three events *A*, *B*, *C*. We say that they are **independent** if any two are independent:

 $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B) \qquad \mathbb{P}(A \cap C) = \mathbb{P}(A)\mathbb{P}(C) \qquad \mathbb{P}(B \cap C) = \mathbb{P}(B)\mathbb{P}(C)$

and in addition

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C)$.

Warning!

Neither of the two sets of conditions imply the other.

The reason we impose all conditions is that independence means also that no two events occurring can influence the chance of a third event occurring.

・ロン ・雪 と ・ ヨ と ・ ヨ と

Consider again the • example with the numbered balls, but now with replacement.

・ロン ・四 と ・ ヨ と ・ ヨ と

ъ

Consider again the \leftarrow example with the numbered balls, but now with replacement. The events A, B and C are now independent:

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C) = \frac{6}{15} \times \frac{4}{15} \times \frac{5}{15} = \frac{8}{225} \simeq 3.6\%$

・聞き ・ヨチ ・ヨチー

э.

Consider again the example with the numbered balls, but now with replacement. The events A, B and C are now independent:

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C) = \frac{6}{15} \times \frac{4}{15} \times \frac{5}{15} = \frac{8}{225} \simeq 3.6\%$

Example

What is the probability of getting at least one 🖸 in two rolls of a fair die?

Consider again the example with the numbered balls, but now with replacement. The events A, B and C are now independent:

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C) = \frac{6}{15} \times \frac{4}{15} \times \frac{5}{15} = \frac{8}{225} \simeq 3.6\%$

Example

What is the probability of getting at least one : in two rolls of a fair die? Let A_i be the event of getting : in the ith roll. We are after $\mathbb{P}(A_1 \cup A_2)$.

Consider again the example with the numbered balls, but now with replacement. The events A, B and C are now independent:

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C) = \frac{6}{15} \times \frac{4}{15} \times \frac{5}{15} = \frac{8}{225} \simeq 3.6\%$

Example

What is the probability of getting at least one \bigcirc in two rolls of a fair die? Let A_i be the event of getting \bigcirc in the ith roll. We are after $\mathbb{P}(A_1 \cup A_2)$. The two events are independent, whence by inclusion-exclusion

$$\mathbb{P}(A_1 \cup A_2) = \mathbb{P}(A_1) + \mathbb{P}(A_2) - \mathbb{P}(A_1 \cap A_2)$$

Consider again the example with the numbered balls, but now with replacement. The events A, B and C are now independent:

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C) = \frac{6}{15} \times \frac{4}{15} \times \frac{5}{15} = \frac{8}{225} \simeq 3.6\%$

Example

What is the probability of getting at least one i in two rolls of a fair die? Let A_i be the event of getting i in the ith roll. We are after $\mathbb{P}(A_1 \cup A_2)$. The two events are independent, whence by inclusion-exclusion

$$\mathbb{P}(A_1 \cup A_2) = \mathbb{P}(A_1) + \mathbb{P}(A_2) - \mathbb{P}(A_1)\mathbb{P}(A_2)$$

Consider again the example with the numbered balls, but now with replacement. The events A, B and C are now independent:

 $\mathbb{P}(A \cap B \cap C) = \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C) = \frac{6}{15} \times \frac{4}{15} \times \frac{5}{15} = \frac{8}{225} \simeq 3.6\%$

Example

What is the probability of getting at least one \bigcirc in two rolls of a fair die? Let A_i be the event of getting \bigcirc in the ith roll. We are after $\mathbb{P}(A_1 \cup A_2)$. The two events are independent, whence by inclusion-exclusion

$$\mathbb{P}(A_1 \cup A_2) = \mathbb{P}(A_1) + \mathbb{P}(A_2) - \mathbb{P}(A_1)\mathbb{P}(A_2)$$
$$= \frac{1}{6} + \frac{1}{6} - \frac{1}{36} = \frac{11}{36}.$$

A fair coin is tossed 10 times. What is the probability of getting all heads? of getting HTHTHTHT? of getting TTTTTTTTH?

A fair coin is tossed 10 times. What is the probability of getting all heads? of getting HTHTHTHTHT? of getting TTTTTTTTH? Getting H or T in one toss of the coin is independent from getting H or T on a different toss.

くぼう くほう くほう

A fair coin is tossed 10 times. What is the probability of getting all heads? of getting HTHTHTHTHT? of getting TTTTTTTTH? Getting H or T in one toss of the coin is independent from getting H or T on a different toss. Since the coin if fair, $\mathbb{P}(H) = \mathbb{P}(T) = \frac{1}{2}$, whence the probability of getting any one desired outcome after 10 tosses is $(\frac{1}{2})^{10}$.

伺い イヨン イヨン

A fair coin is tossed 10 times. What is the probability of getting all heads? of getting HTHTHTHTH? of getting TTTTTTTTH? Getting H or T in one toss of the coin is independent from getting H or T on a different toss. Since the coin if fair, $\mathbb{P}(H) = \mathbb{P}(T) = \frac{1}{2}$, whence the probability of getting any one desired outcome after 10 tosses is $(\frac{1}{2})^{10}$. What about if the coin is not fair?

A D A D A D A

20/23

A fair coin is tossed 10 times. What is the probability of getting all heads? of getting HTHTHTHTHT? of getting TTTTTTTTH? Getting H or T in one toss of the coin is independent from getting H or T on a different toss. Since the coin if fair, $\mathbb{P}(H) = \mathbb{P}(T) = \frac{1}{2}$, whence the probability of getting any one desired outcome after 10 tosses is $(\frac{1}{2})^{10}$. What about if the coin is not fair? Let $\mathbb{P}(H) = p$ and $\mathbb{P}(T) = q = 1 - p$.

伺下 イヨト イヨト

A fair coin is tossed 10 times. What is the probability of getting all heads? of getting HTHTHTHTHT? of getting TTTTTTTTH? Getting H or T in one toss of the coin is independent from getting H or T on a different toss. Since the coin if fair, $\mathbb{P}(H) = \mathbb{P}(T) = \frac{1}{2}$, whence the probability of getting any one desired outcome after 10 tosses is $(\frac{1}{2})^{10}$. What about if the coin is not fair? Let $\mathbb{P}(H) = p$ and $\mathbb{P}(T) = q = 1 - p$. Then by independence,

$$\begin{split} \mathbb{P}(HHHHHHHHH) &= \mathbb{P}(H)^{10} = p^{10} \\ \mathbb{P}(HTHTHTHTHT) &= \mathbb{P}(H)\mathbb{P}(T)\mathbb{P}(H) \dots \mathbb{P}(T) = p^5(1-p)^5 \\ \mathbb{P}(TTTTTTTTTH) &= \mathbb{P}(T)^9\mathbb{P}(H) = (1-p)^9p \end{split}$$

くぼう くほう くほう

Independence vs. pairwise independence

Definition

A family A_1, A_2, \ldots, A_n of events is **pairwise independent** if

 $\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i)\mathbb{P}(A_j) \qquad \forall \quad 1 \leqslant i < j \leqslant n$

Independence vs. pairwise independence

Definition

A family A_1, A_2, \ldots, A_n of events is **pairwise independent** if

 $\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i)\mathbb{P}(A_j) \qquad \forall \quad 1 \leqslant i < j \leqslant n$

and it is **independent** if, for any $1 \leq j_1 < j_2 < \cdots < j_k \leq n$,

 $\mathbb{P}(A_{j_1} \cap A_{j_2} \cap \dots \cap A_{j_k}) = \mathbb{P}(A_{j_1})\mathbb{P}(A_{j_2}) \cdots \mathbb{P}(A_{j_k})$

< □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □ > < □

Independence vs. pairwise independence

Definition

A family A_1, A_2, \ldots, A_n of events is **pairwise independent** if

 $\mathbb{P}(A_i \cap A_j) = \mathbb{P}(A_i)\mathbb{P}(A_j) \qquad \forall \quad 1 \leqslant i < j \leqslant n$

and it is **independent** if, for any $1 \leq j_1 < j_2 < \cdots < j_k \leq n$,

 $\mathbb{P}(A_{j_1} \cap A_{j_2} \cap \dots \cap A_{j_k}) = \mathbb{P}(A_{j_1})\mathbb{P}(A_{j_2}) \cdots \mathbb{P}(A_{j_k})$

Whereas independence implies pairwise independence, the converse **is not true**.

Two fair coins are tossed. Consider the events

- A = "the first toss is a head"
- B = "the second toss is a head"
- $C = A \triangle B$ = "exactly one toss is a head"

Two fair coins are tossed. Consider the events

- A = "the first toss is a head"
- B = "the second toss is a head"
- $C = A \triangle B$ = "exactly one toss is a head"

Then $\mathbb{P}(A) = \mathbb{P}(B) = \mathbb{P}(C) = \frac{1}{2}$ and

 $\mathbb{P}(A \cap B) = \frac{1}{4} = \mathbb{P}(A)\mathbb{P}(B)$ $\mathbb{P}(A \cap C) = \frac{1}{4} = \mathbb{P}(A)\mathbb{P}(C)$ $\mathbb{P}(B \cap C) = \frac{1}{4} = \mathbb{P}(B)\mathbb{P}(C) ,$

whence the events are pairwise independent.

< 🗇 > < 🖻 > < 🖻 >

Two fair coins are tossed. Consider the events

- A = "the first toss is a head"
- B = "the second toss is a head"
- $C = A \triangle B$ = "exactly one toss is a head"

Then $\mathbb{P}(A) = \mathbb{P}(B) = \mathbb{P}(C) = \frac{1}{2}$ and

 $\mathbb{P}(A \cap B) = \frac{1}{4} = \mathbb{P}(A)\mathbb{P}(B)$ $\mathbb{P}(A \cap C) = \frac{1}{4} = \mathbb{P}(A)\mathbb{P}(C)$ $\mathbb{P}(B \cap C) = \frac{1}{4} = \mathbb{P}(B)\mathbb{P}(C) ,$

whence the events are pairwise independent. However

 $\mathbb{P}(A \cap B \cap C) = \boldsymbol{0} \neq \mathbb{P}(A)\mathbb{P}(B)\mathbb{P}(C)$,

and the events are **not** independent.

A > A > A > A >

The conditional probability of A occuring given that B occurs is

 $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$

ヘロト ヘヨト ヘヨト ヘヨト

э.

The conditional probability of A occuring given that B occurs is

 $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$

• Conditional probability is a probability

イロン イロン イヨン イヨン

э.

The conditional probability of A occuring given that B occurs is

 $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$

- Conditional probability is a probability
- Multiplication rules:

 $\mathbb{P}(A_1 \cap A_2) = \mathbb{P}(A_2|A_1)\mathbb{P}(A_1)$ $\mathbb{P}(A_1 \cap A_2 \cap A_3) = \mathbb{P}(A_3|A_1 \cap A_2)\mathbb{P}(A_2|A_1)\mathbb{P}(A_1)$

The conditional probability of A occuring given that B occurs is

 $\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$

- Conditional probability is a probability
- Multiplication rules:

 $\mathbb{P}(A_1 \cap A_2) = \mathbb{P}(A_2|A_1)\mathbb{P}(A_1)$ $\mathbb{P}(A_1 \cap A_2 \cap A_3) = \mathbb{P}(A_3|A_1 \cap A_2)\mathbb{P}(A_2|A_1)\mathbb{P}(A_1)$

• Events A, B are independent if $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$

The conditional probability of A occuring given that B occurs is

$$\mathbb{P}(A|B) = \frac{\mathbb{P}(A \cap B)}{\mathbb{P}(B)}$$

- Conditional probability is a probability
- Multiplication rules:

 $\mathbb{P}(A_1 \cap A_2) = \mathbb{P}(A_2|A_1)\mathbb{P}(A_1)$ $\mathbb{P}(A_1 \cap A_2 \cap A_3) = \mathbb{P}(A_3|A_1 \cap A_2)\mathbb{P}(A_2|A_1)\mathbb{P}(A_1)$

- Events A, B are independent if $\mathbb{P}(A \cap B) = \mathbb{P}(A)\mathbb{P}(B)$
- More generally, a countable family $\{A_i\}, i \in I$, of events are independent if

$$\mathbb{P}(\bigcap_{j\in J}A_j)=\prod_{j\in J}\mathbb{P}(A_j)$$

for every finite subset $J\subseteq \mathrm{I}$