Smoothness Selection

Simon Wood

Mathematical Sciences, University of Bath, U.K.

Smoothness selection approaches

- The smoothing model y_i = f(x_i) + ε_i, ε_i ~ N(0, σ²), is represented via a basis expansion of f, with coefficients β.
- The β estimates are β̂ = arg min_β ||**y** − **X**β||² + λβ^T**S**β where **X** is the model matrix derived from the basis, and **S** is the wiggliness penalty matrix.
- λ controls smoothness how should it be chosen?
- There are 3 main statistical approaches
 - 1. Choose λ to minimize error in predicting new data.
 - 2. Treat smooths as random effects, following the Bayesian smoothing model, and estimate λ as a variance parameter using a marginal likelihood approach.
 - 3. Go fully Bayesian by completing the Bayesian model with a prior on λ (requires simulation and not pursued here).

Prediction error: C_p /UBRE

- Suppose σ^2 is known, and let $\mathbf{A} = \mathbf{X} (\mathbf{X}^{\mathsf{T}} \mathbf{X} + \lambda \mathbf{S})^{-1} \mathbf{X}^{\mathsf{T}}$.
- $\hat{\mu} = \mathsf{A}\mathsf{y}$ where $\mathbb{E}(\mathsf{y}) = \mu$, so consider

$$\begin{aligned} \|\boldsymbol{\mu} - \hat{\boldsymbol{\mu}}\|^2 &= \|\boldsymbol{\mu} - \mathbf{A}\mathbf{y}\|^2 = \|\mathbf{y} - \mathbf{A}\mathbf{y} - \boldsymbol{\epsilon}\|^2 \\ &= \|\mathbf{y} - \mathbf{A}\mathbf{y}\|^2 + \boldsymbol{\epsilon}^{\mathsf{T}}\boldsymbol{\epsilon} - 2\boldsymbol{\epsilon}^{\mathsf{T}}(\mathbf{y} - \mathbf{A}\mathbf{y}) \\ &= \|\mathbf{y} - \mathbf{A}\mathbf{y}\|^2 + \boldsymbol{\epsilon}^{\mathsf{T}}\boldsymbol{\epsilon} - 2\boldsymbol{\epsilon}^{\mathsf{T}}(\boldsymbol{\mu} + \boldsymbol{\epsilon}) + 2\boldsymbol{\epsilon}^{\mathsf{T}}\mathbf{A}(\boldsymbol{\mu} + \boldsymbol{\epsilon}) \\ &= \|\mathbf{y} - \mathbf{A}\mathbf{y}\|^2 - \boldsymbol{\epsilon}^{\mathsf{T}}\boldsymbol{\epsilon} - 2\boldsymbol{\epsilon}^{\mathsf{T}}\boldsymbol{\mu} + 2\boldsymbol{\epsilon}^{\mathsf{T}}\mathbf{A}\boldsymbol{\mu} + 2\boldsymbol{\epsilon}^{\mathsf{T}}\mathbf{A}\boldsymbol{\epsilon} \end{aligned}$$

- ► Hence $\mathbb{E} \| \boldsymbol{\mu} \hat{\boldsymbol{\mu}} \|^2 = \mathbb{E} \| \mathbf{y} \mathbf{A} \mathbf{y} \|^2 n\sigma^2 + 2\sigma^2 \operatorname{tr}(\mathbf{A})$
- Estimating $\mathbb{E} \| \mathbf{y} \mathbf{A} \mathbf{y} \|^2$ yields . . .

$$\mathcal{C}_{p} = \|\mathbf{y} - \mathbf{A}\mathbf{y}\|^{2} - n\sigma^{2} + 2\sigma^{2} \operatorname{tr}(\mathbf{A})$$

• Can choose λ to minimize C_p .

σ^2 unknown: cross validation



1. Choose λ to try to minimize the error predicting new data.

- 2. Minimize the average error in predicting single datapoints *omitted* from the fit. Each datum left out once in average.
- 3. It turns out that

$$\mathcal{V}_o(\lambda) = \frac{1}{n} \sum_i (y_i - \hat{\mu}_i^{[-i]})^2 = \frac{1}{n} \sum_i \frac{(y_i - \hat{\mu}_i)^2}{(1 - A_{ii})^2}$$

OCV not invariant



 OCV is not invariant in an odd way. If Q is orthogonal then fitting objective

$$\|\mathbf{Q}\mathbf{y} - \mathbf{Q}\mathbf{X}\boldsymbol{\beta}\|^2 + \lambda \boldsymbol{\beta}^\mathsf{T}\mathbf{S}\boldsymbol{\beta}$$

yields identical inferences about β as the original objective, but it gives a different \mathcal{V}_o .

GCV: generalized cross validation

If we find the Q that causes the leading diagonal elements of A to be constant, and then perform OCV, the result is the invariant alternative GCV:

$$\mathcal{V}_g = rac{n \|\mathbf{y} - \hat{\boldsymbol{\mu}}\|^2}{\{n - \operatorname{tr}(\mathbf{A})\}^2}$$

- It is easy to show that tr(A) = tr(F), where F is the degrees of freedom matrix.
- In addition to invariance, GCV is much easier to optimize efficiently in the multiple smoothing parameter case.

REML/ML λ estimation

The Bayesian smooth model is

$$\mathbf{y} = \mathbf{X} \boldsymbol{\beta} + \boldsymbol{\epsilon}, \ \ \boldsymbol{\beta} \sim N(\mathbf{0}, \mathbf{S}^{-} \sigma^{2} / \lambda), \ \ \boldsymbol{\epsilon} \sim N(\mathbf{0}, \mathbf{I} \sigma^{2})$$

- This can be viewed as a mixed model for computational purposes, but the impropiety of f(β) is awkward.
- To fix this, find the eigen-decomposition $\mathbf{S} = \mathbf{U} \mathbf{\Lambda} \mathbf{U}^{\mathsf{T}}$
- Reparameterize β' = U^Tβ and let Λ₊ denote the diagonal matrix of +ve eigenvalues.
- ► Now $\beta^{\mathsf{T}} \mathbf{S} \beta = \beta'^{\mathsf{T}} \Lambda \beta' = \mathbf{b}^{\mathsf{T}} \Lambda_{+} \mathbf{b}$ where $\beta' = (\mathbf{b}^{\mathsf{T}}, \gamma^{\mathsf{T}})^{\mathsf{T}}$.
- ▶ Now partition $\mathbf{X}' = \mathbf{X}\mathbf{U} = (\mathbf{Z} : \tilde{\mathbf{X}})$, so that the model becomes

$$\mathbf{y} = \mathbf{ ilde{X}} \boldsymbol{\gamma} + \mathbf{Z} \mathbf{b} + \boldsymbol{\epsilon}, \;\; \mathbf{b} \sim \mathcal{N}(\mathbf{0}, \mathbf{\Lambda}_+^{-1} \sigma^2 / \lambda), \;\; \boldsymbol{\epsilon} \sim \mathcal{N}(\mathbf{0}, \mathbf{I} \sigma^2)$$

$\mathsf{REML}/\mathsf{ML}$ λ estimation

- ► Now that the model is in standard mixed model form, mixed model methods can estimate *λ* as a variance parameter.
- ▶ MLE or REML can be used.
- From a Bayesian perspective we are being empirical Bayesians and using marginal likelihood.
- Notice that the restricted/marginal likelihood has the form

$$\int f(\mathbf{y}|\boldsymbol{eta})f(\boldsymbol{eta})d\boldsymbol{eta}$$

- That is, we are taking the expectation of the likelihood over the prior on β.
- From this perspective it is possible to plot why the approach is intuitively sensible.

Basic principle of ML smoothness selection



- 1. Choose λ to maximize the average likelihood of random draws from the prior implied by λ .
- 2. If λ too low, then almost all draws are too variable to have high likelihood. If λ too high, then draws all underfit and have low likelihood. The right λ maximizes the proportion of draws close enough to data to give high likelihood.
- 3. Formally, maximize e.g. $\mathcal{V}_r(\lambda) = \log \int f(\mathbf{y}|\boldsymbol{\beta}) f_{\lambda}(\boldsymbol{\beta}) d\boldsymbol{\beta}$.

Prediction error vs. likelihood λ estimation



- 1. Pictures show GCV and REML scores for different replicates from same truth.
- 2. Compared to REML, GCV penalizes overfit only weakly, and so tends to undersmooth.

Are smoothers *really* random effects?

- Most times that smooth functions are used in models, the modeller believes that the function is a fixed state of nature.
- i.e. the assumption is that the true function is something that would stay fixed on replication of the dataset.
- So we are *really* being Bayesian in treating the function as random.
- If the function was a true frequentist random effect then we would expect to get a different random draw from its prior at each dataset replication. This almost never makes sense.
- Does this mean that using mixed modelling methods is wrong?
- No. It just happens that the mixed model methods can conveniently compute the Bayesian answers for us.